

Hamiltonian Monte Carlo & eigenvector continuation for *ab initio* nuclear physics

Andreas Ekström



CHALMERS
UNIVERSITY OF TECHNOLOGY



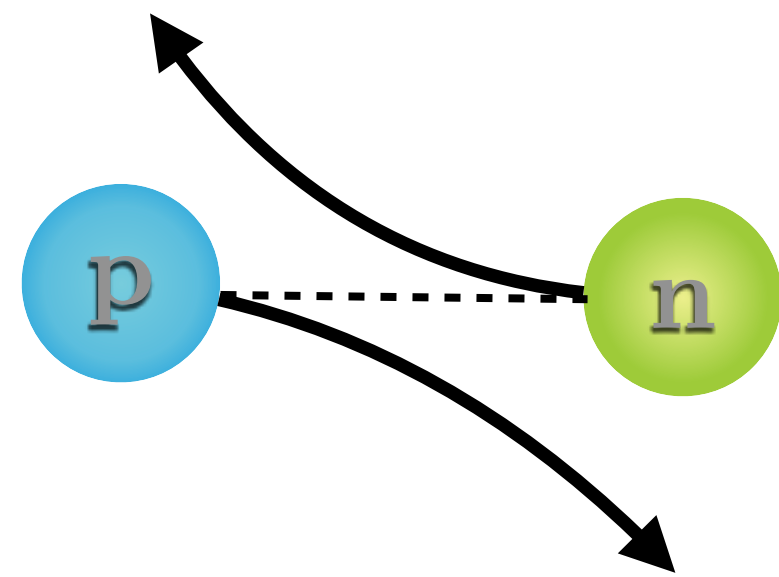
Swedish
Research Council



Contents

Ab initio modelling of nuclear systems

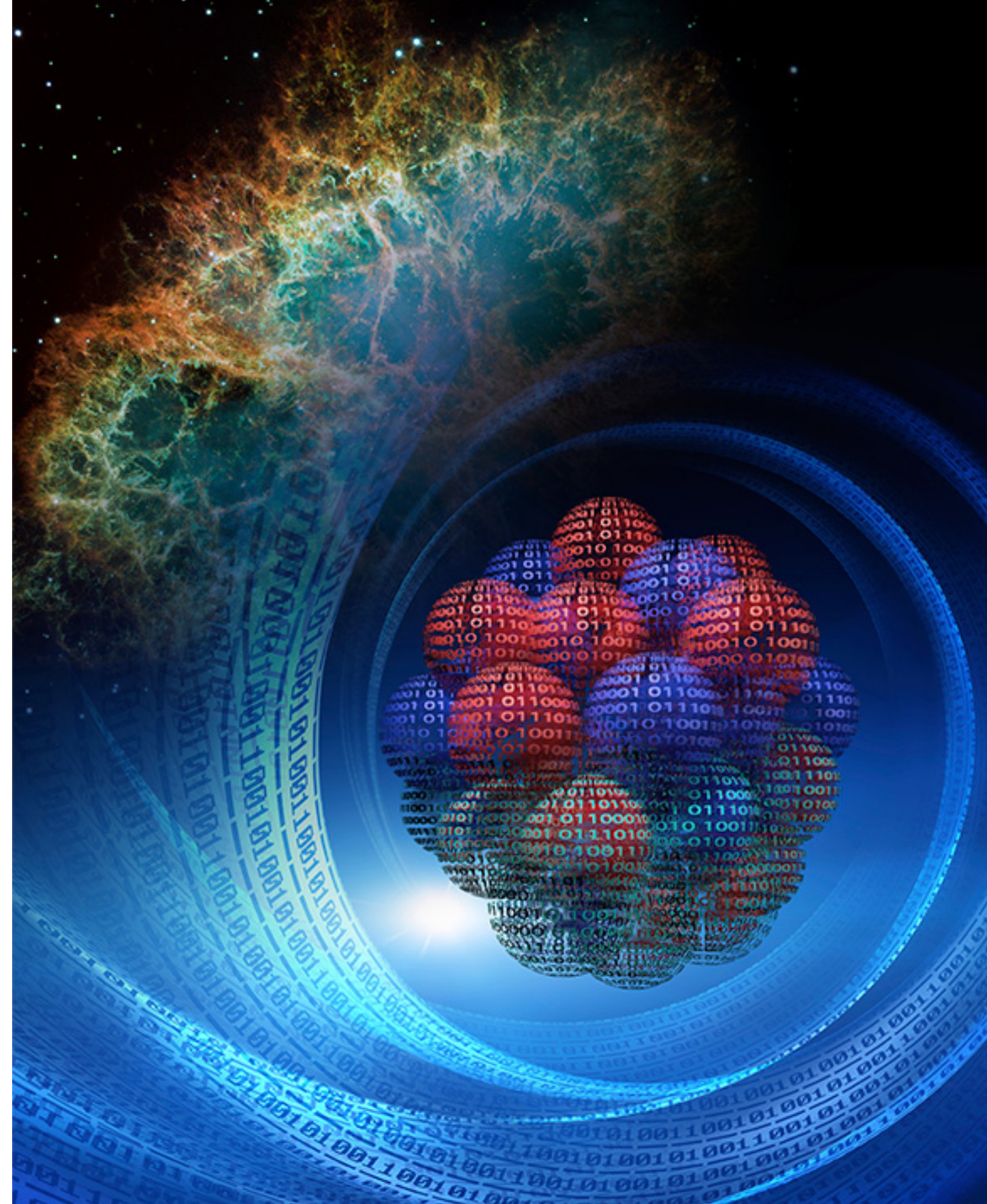
Bayesian inference of neutron-proton scattering



a prototypical system for studying the
strong nuclear interaction

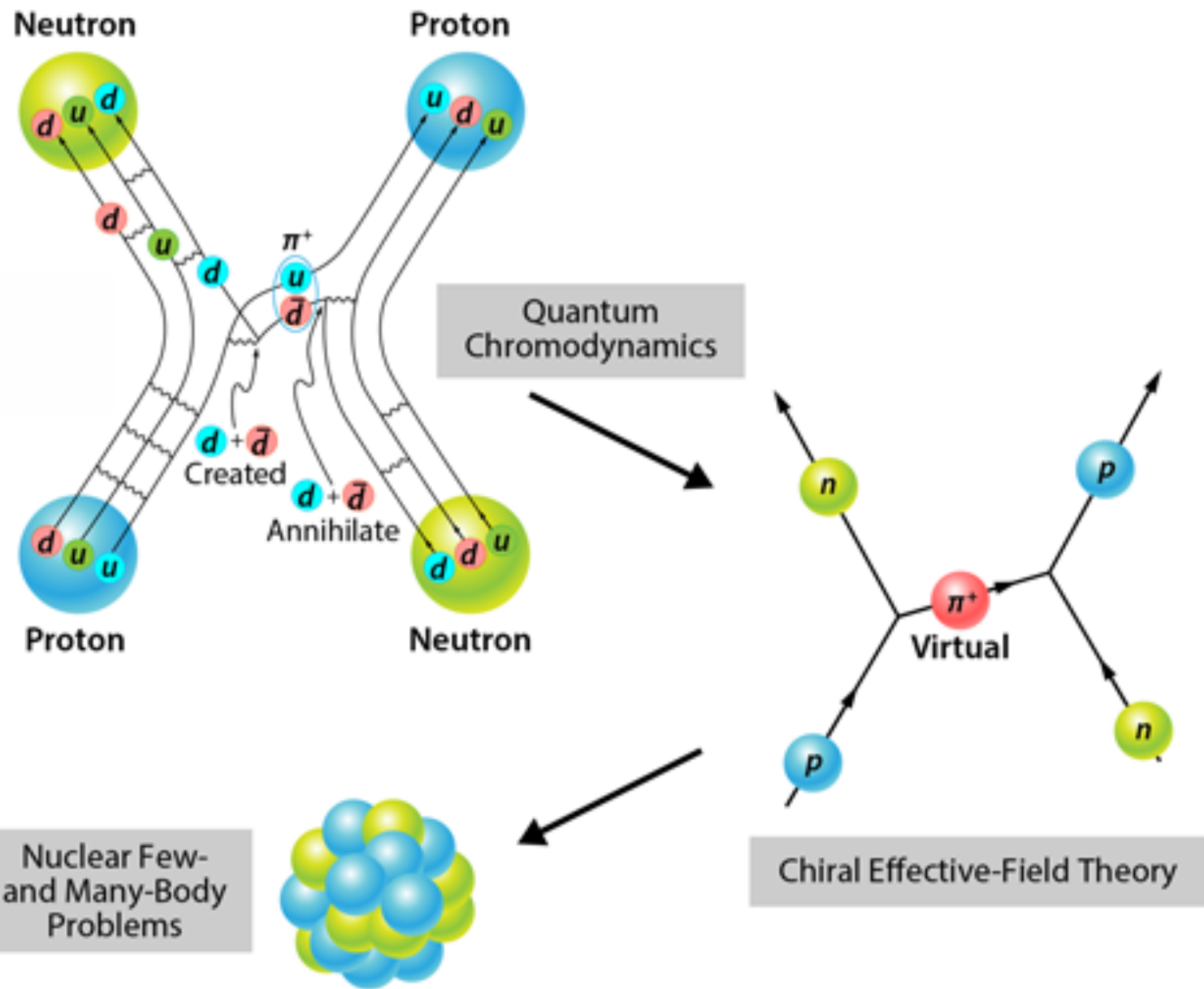
For details, see our recent preprint

Isak Svensson, Andreas Ekström, Christian Forssén arXiv:2304.02004 [nucl-th]



Ab initio modelling of nuclear systems

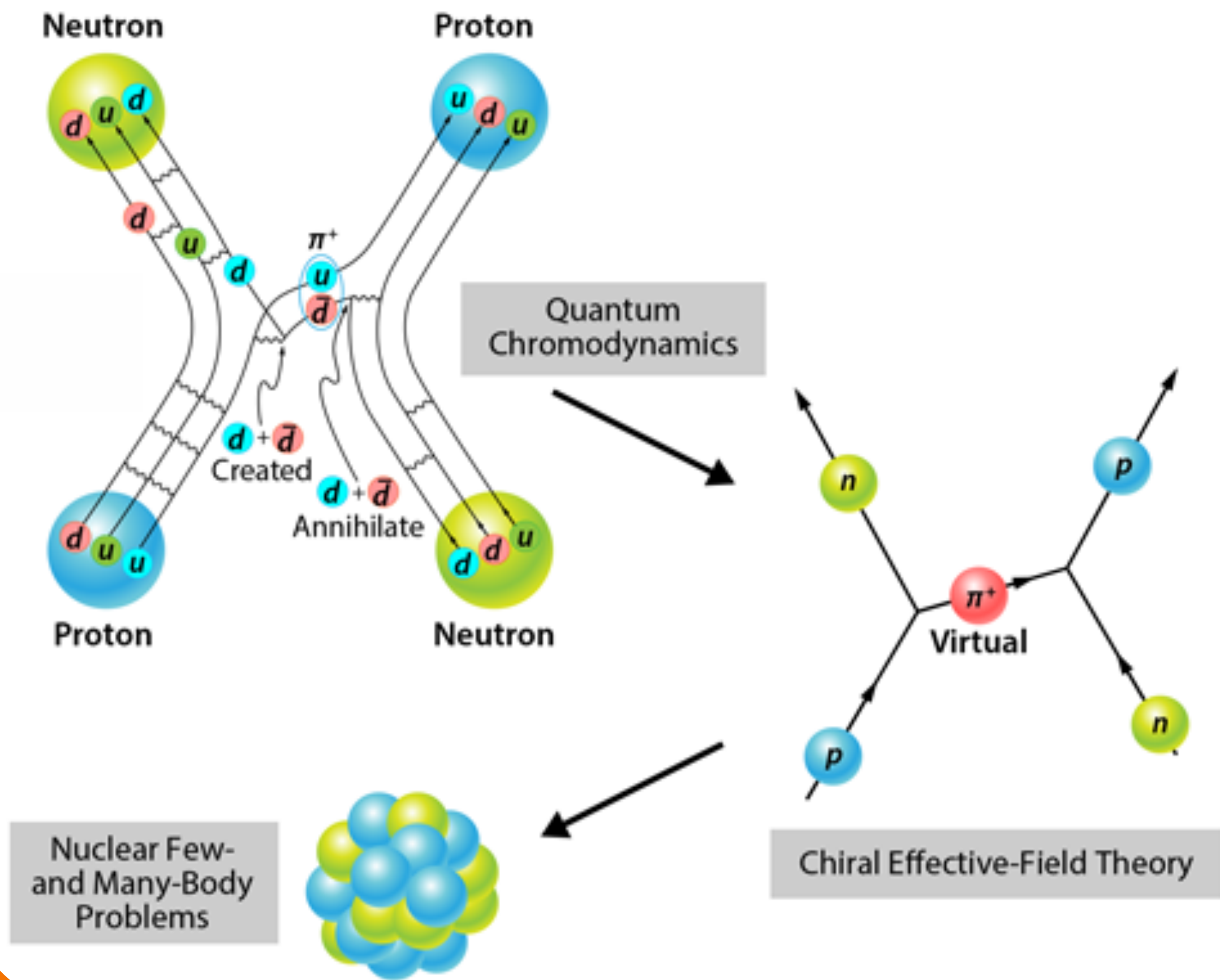
The Idea



credit: APS/Alan Stonebraker

Ab initio modelling of nuclear systems

The Idea

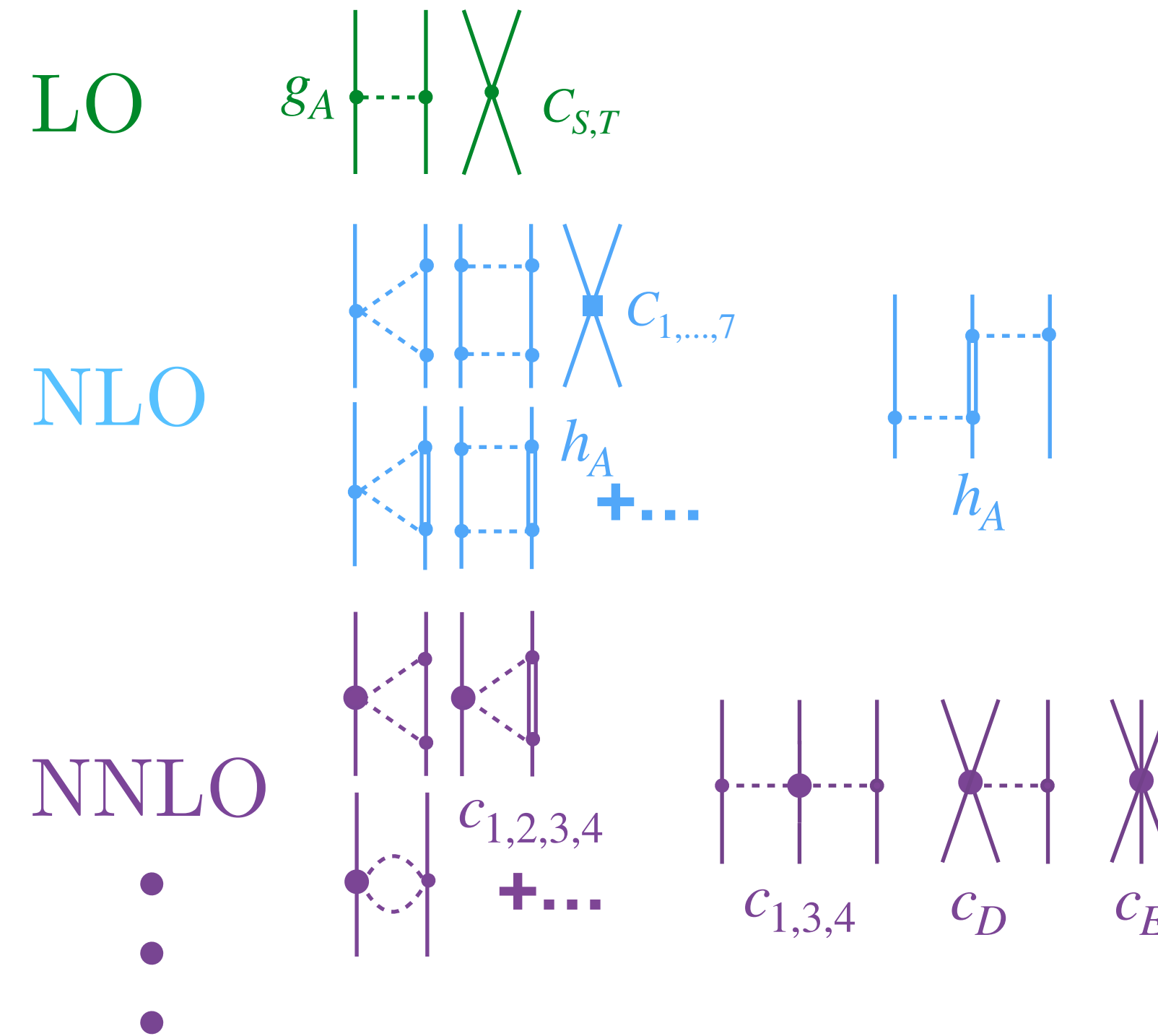


credit: APS/Alan Stonebraker

The Model

$$\hat{H} |\Psi\rangle = E |\Psi\rangle$$

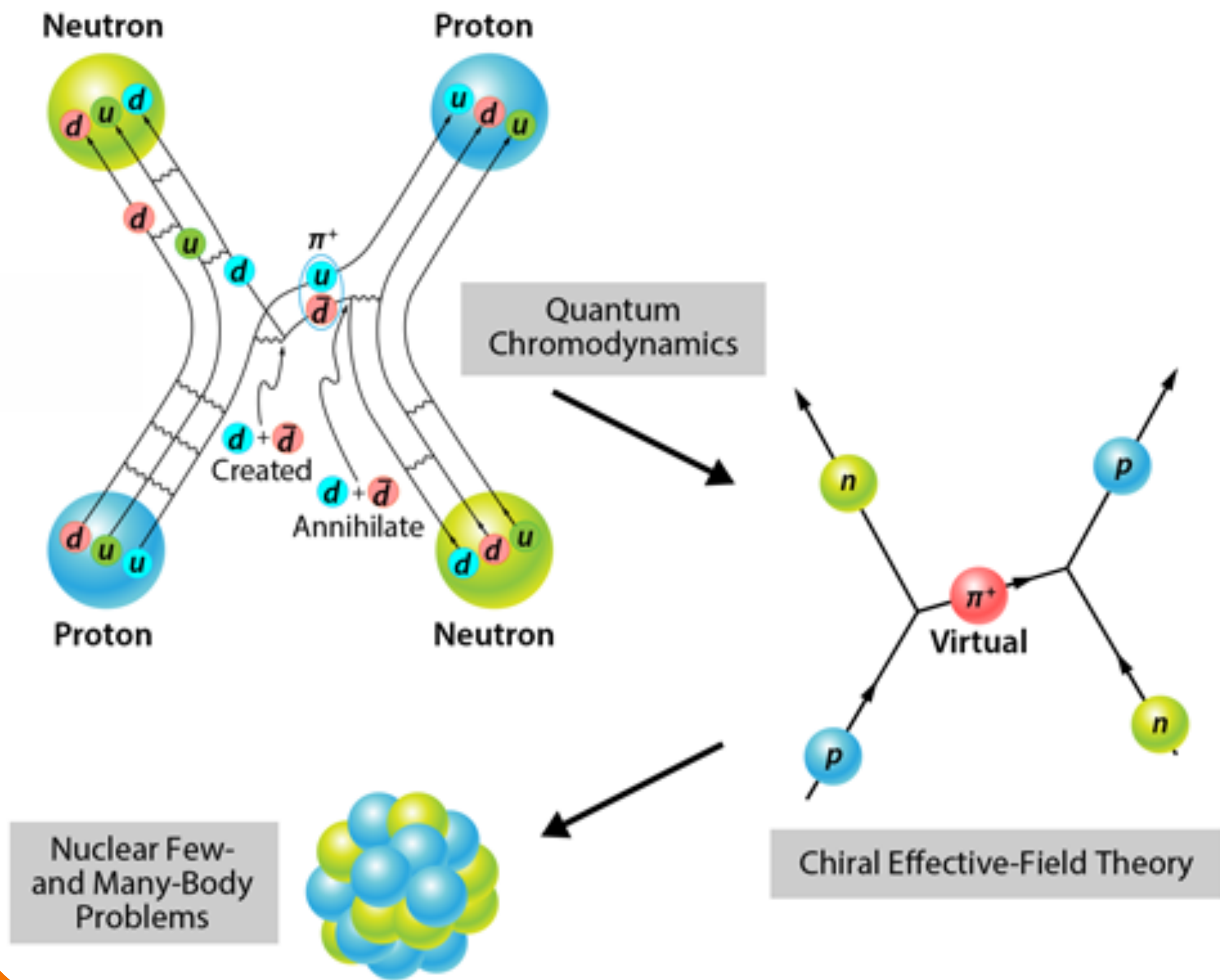
$$\hat{H}(\vec{\alpha}) = \hat{T} + \hat{V}(\vec{\alpha})$$



Weinberg, van Kolck, Meißner, Epelbaum, Machleidt, Entem, ...

Ab initio modelling of nuclear systems

The Idea



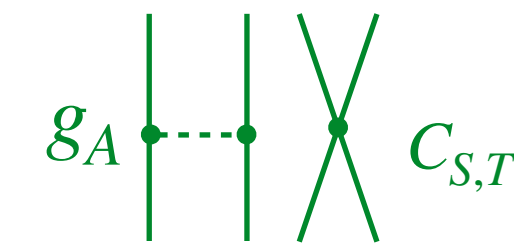
credit: APS/Alan Stonebraker

The Model

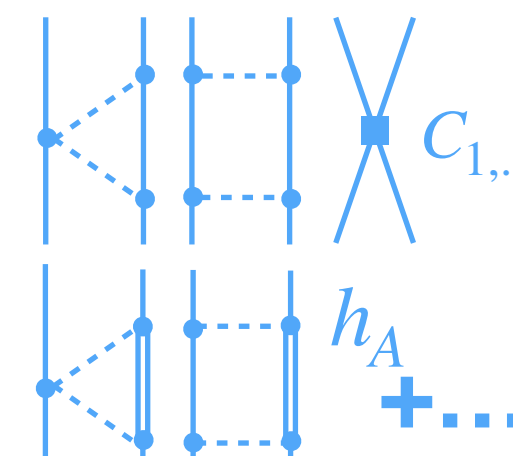
$$\hat{H} |\Psi\rangle = E |\Psi\rangle$$

$$\hat{H}(\vec{\alpha}) = \hat{T} + \hat{V}(\vec{\alpha})$$

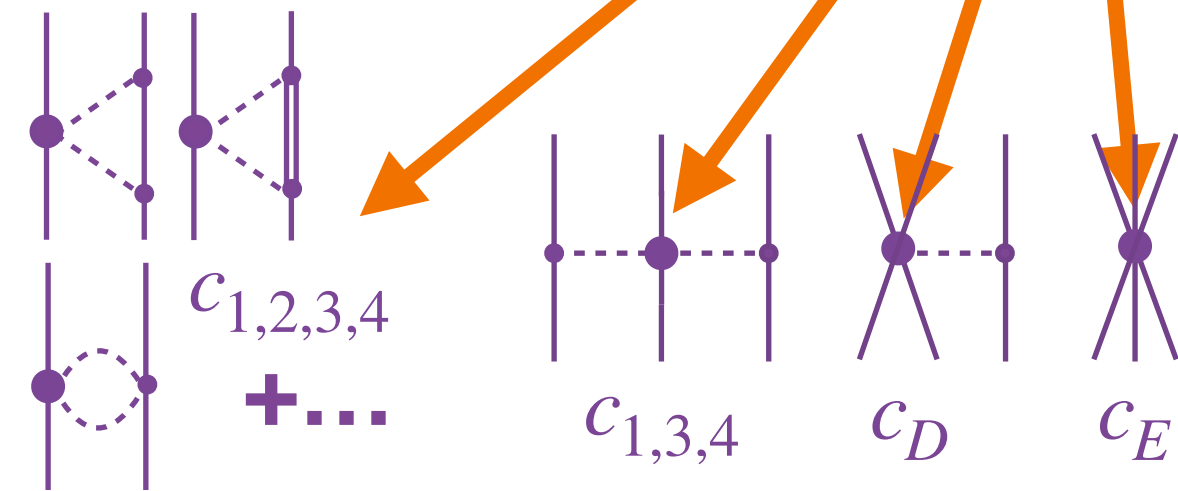
LO



NLO



NNLO



Model parameters

Weinberg, van Kolck, Meißner, Epelbaum, Machleidt, Entem, ...

Why *ab initio* nuclear theory?

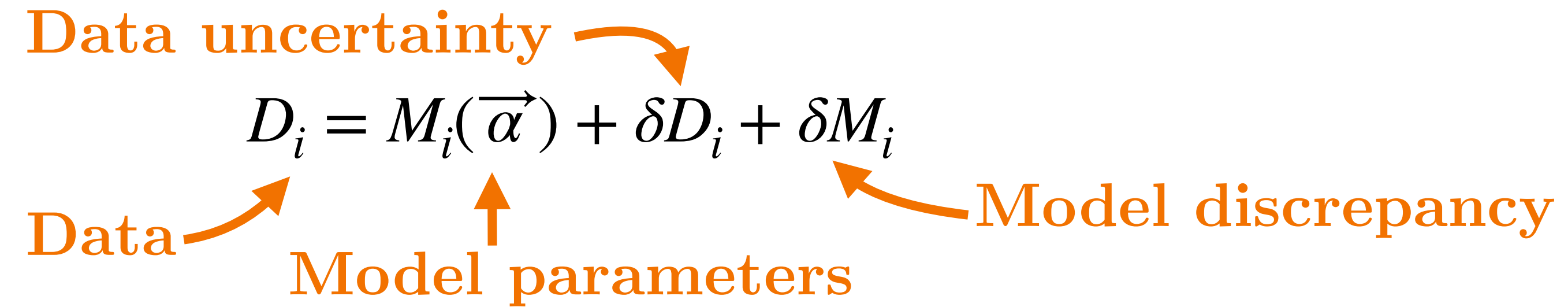
Data uncertainty

$$D_i = M_i(\vec{\alpha}) + \delta D_i + \delta M_i$$

Data

Model parameters

Model discrepancy



Why *ab initio* nuclear theory?

Data uncertainty

$$D_i = M_i(\vec{\alpha}) + \delta D_i + \delta M_i$$

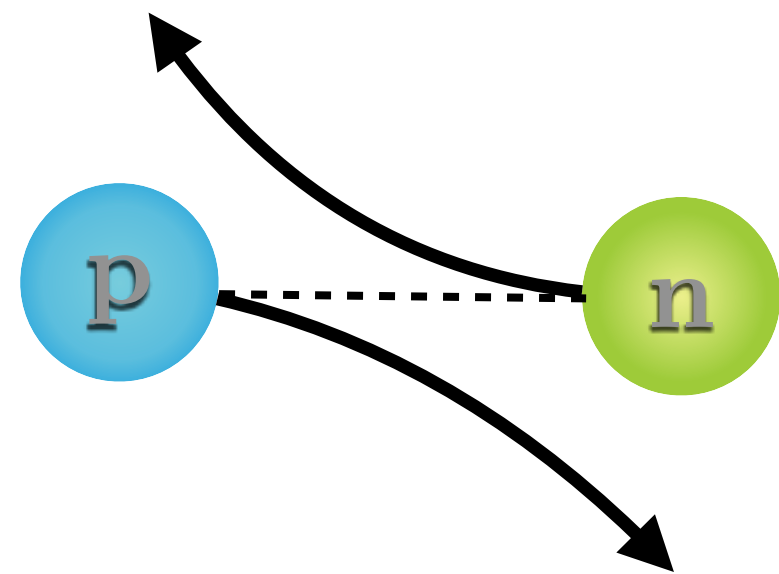
Data

Model parameters

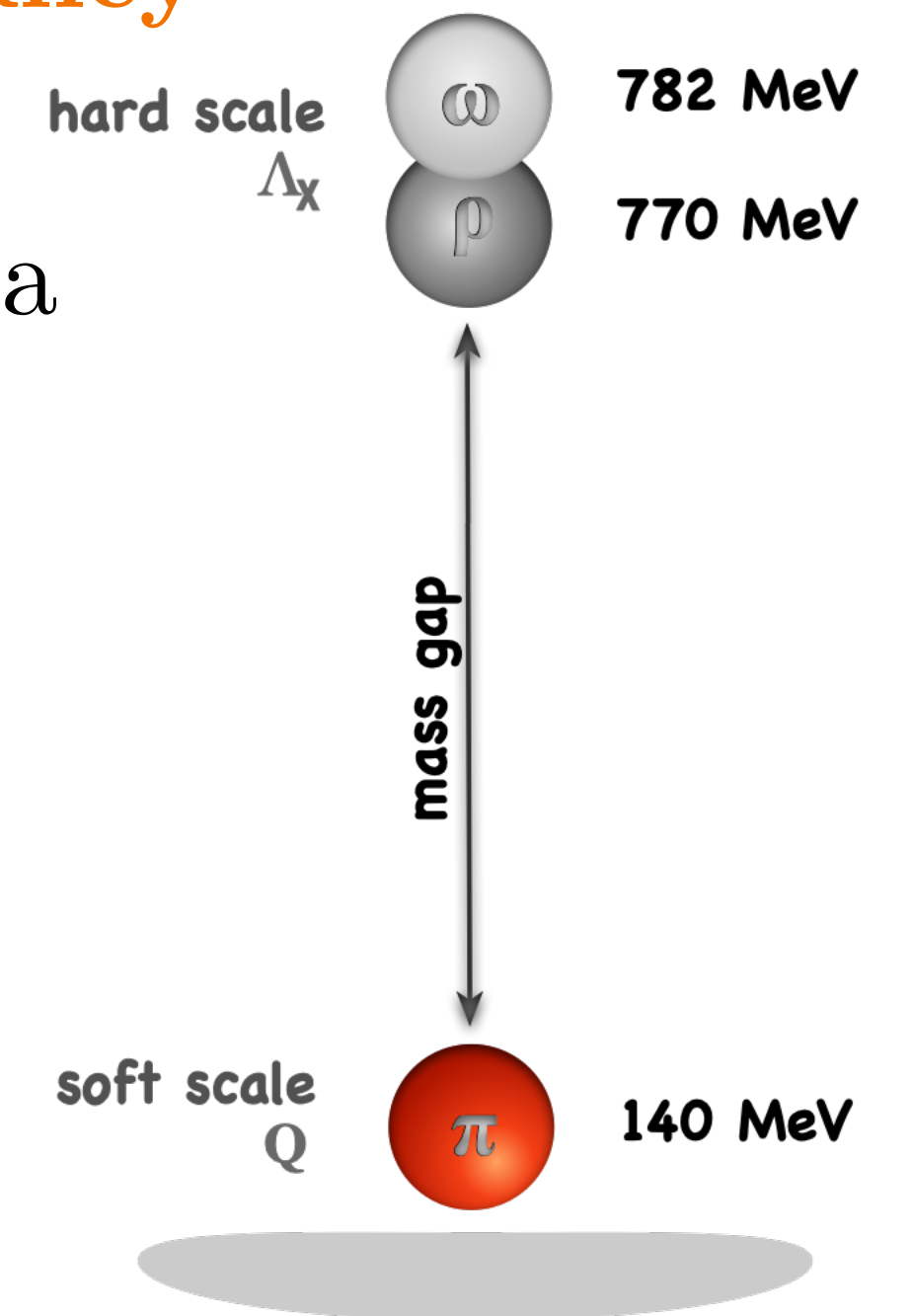
Model discrepancy

We have a model (an effective field theory) for describing data

$$M_i(\vec{\alpha}) = y^{(k)}(\vec{\alpha}; \vec{x}_i) = y_{\text{ref}} \left[\underbrace{\gamma_0 \left(\frac{Q}{\Lambda_\chi} \right)^0}_{\text{LO}} + \underbrace{\gamma_1 \left(\frac{Q}{\Lambda_\chi} \right)^1}_{\text{NLO}} + \underbrace{\gamma_2 \left(\frac{Q}{\Lambda_\chi} \right)^2}_{\text{N2LO}} + \dots + \underbrace{\gamma_k \left(\frac{Q}{\Lambda_\chi} \right)^k}_{\text{NkLO}} \right]$$



nucleon-nucleon
scattering data
(~2000 of relevance)



Why *ab initio* nuclear theory?

Data uncertainty

$$D_i = M_i(\vec{\alpha}) + \delta D_i + \delta M_i$$

Data

Model parameters

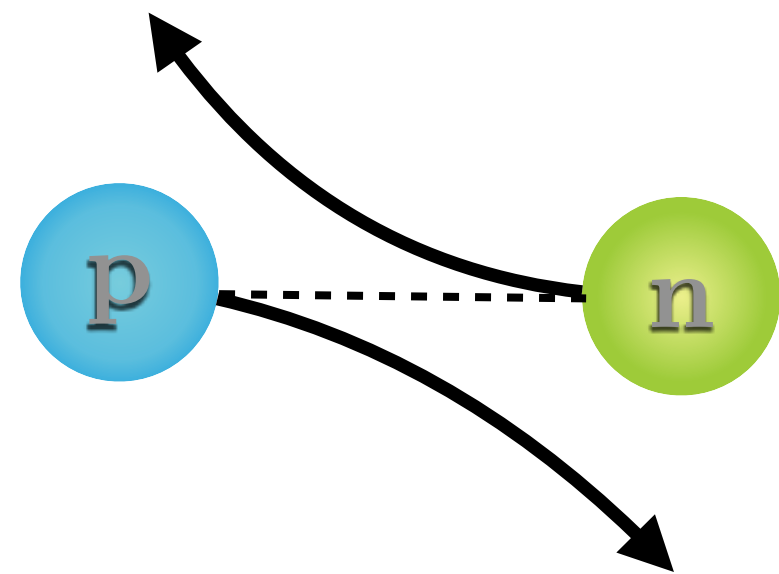
Model discrepancy

We have a model (an effective field theory) for describing data

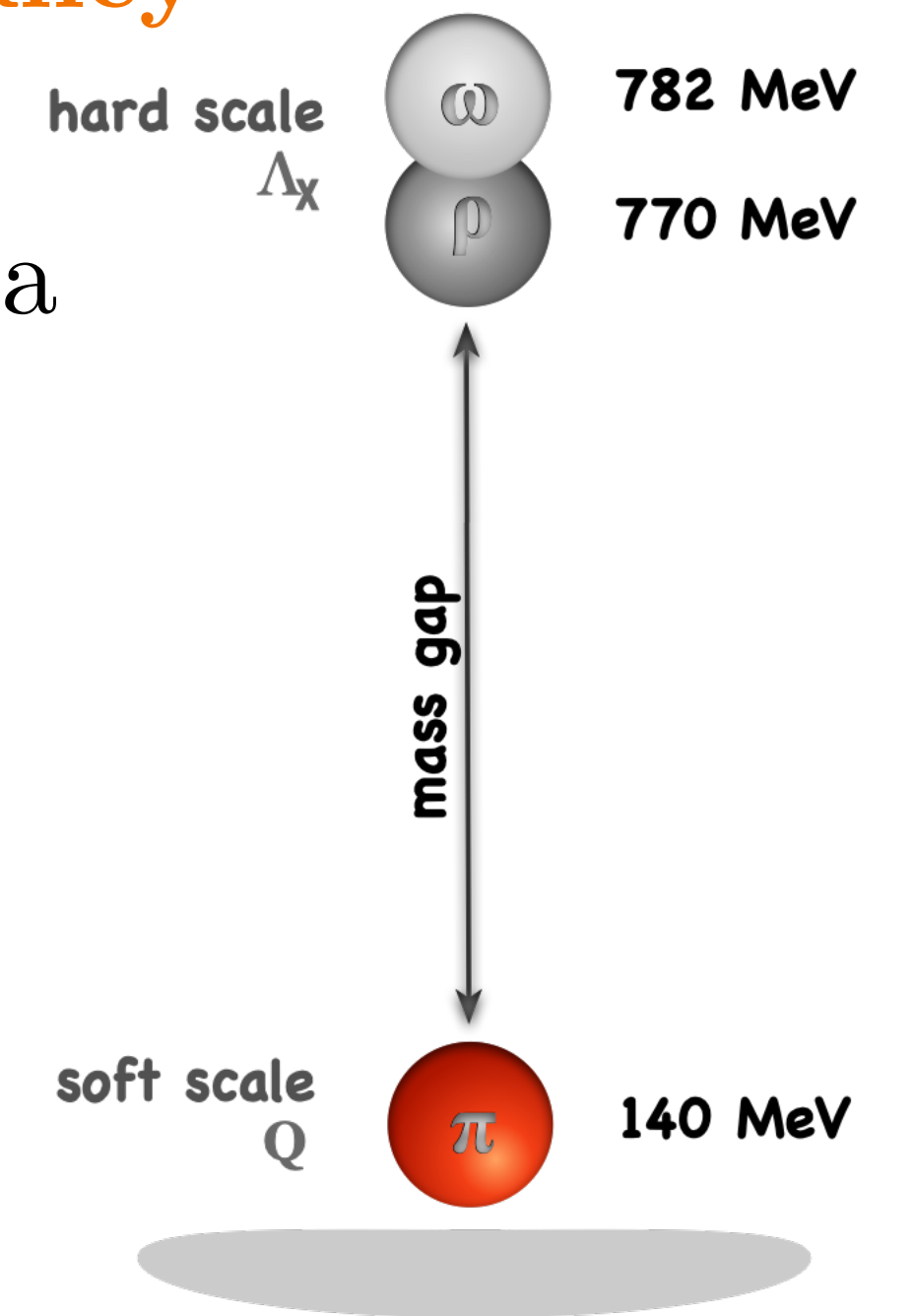
$$M_i(\vec{\alpha}) = y^{(k)}(\vec{\alpha}; \vec{x}_i) = y_{\text{ref}} \left[\underbrace{\gamma_0 \left(\frac{Q}{\Lambda_\chi}\right)^0}_{\text{LO}} + \underbrace{\gamma_1 \left(\frac{Q}{\Lambda_\chi}\right)^1}_{\text{NLO}} + \underbrace{\gamma_2 \left(\frac{Q}{\Lambda_\chi}\right)^2}_{\text{N2LO}} + \dots + \underbrace{\gamma_k \left(\frac{Q}{\Lambda_\chi}\right)^k}_{\text{NkLO}} \right]$$

The systematicity of the *ab initio* method creates an **inferential advantage** since we claim to know something about the model discrepancy.

$$\delta M_i = y_{\text{ref}} \sum_{i=k+1}^{\infty} \gamma_i \left(\frac{Q}{\Lambda_\chi}\right)^i$$



nucleon-nucleon
scattering data
(~2000 of relevance)



Why *ab initio* nuclear theory?

Data uncertainty

$$D_i = M_i(\vec{\alpha}) + \delta D_i + \delta M_i$$

Data

Model parameters

Model discrepancy

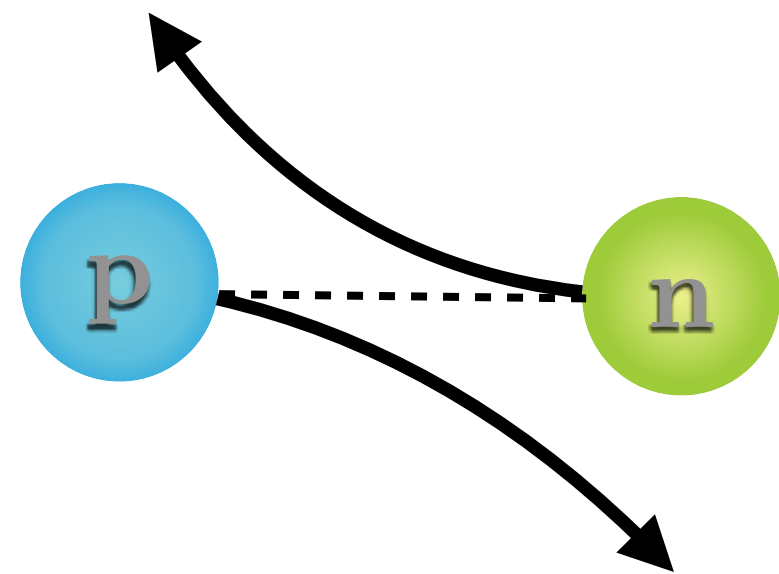
We have a model (an effective field theory) for describing data

$$M_i(\vec{\alpha}) = y^{(k)}(\vec{\alpha}; \vec{x}_i) = y_{\text{ref}} \left[\underbrace{\gamma_0 \left(\frac{Q}{\Lambda_\chi} \right)^0}_{\text{LO}} + \underbrace{\gamma_1 \left(\frac{Q}{\Lambda_\chi} \right)^1}_{\text{NLO}} + \underbrace{\gamma_2 \left(\frac{Q}{\Lambda_\chi} \right)^2}_{\text{N2LO}} + \dots + \underbrace{\gamma_k \left(\frac{Q}{\Lambda_\chi} \right)^k}_{\text{NkLO}} \right]$$

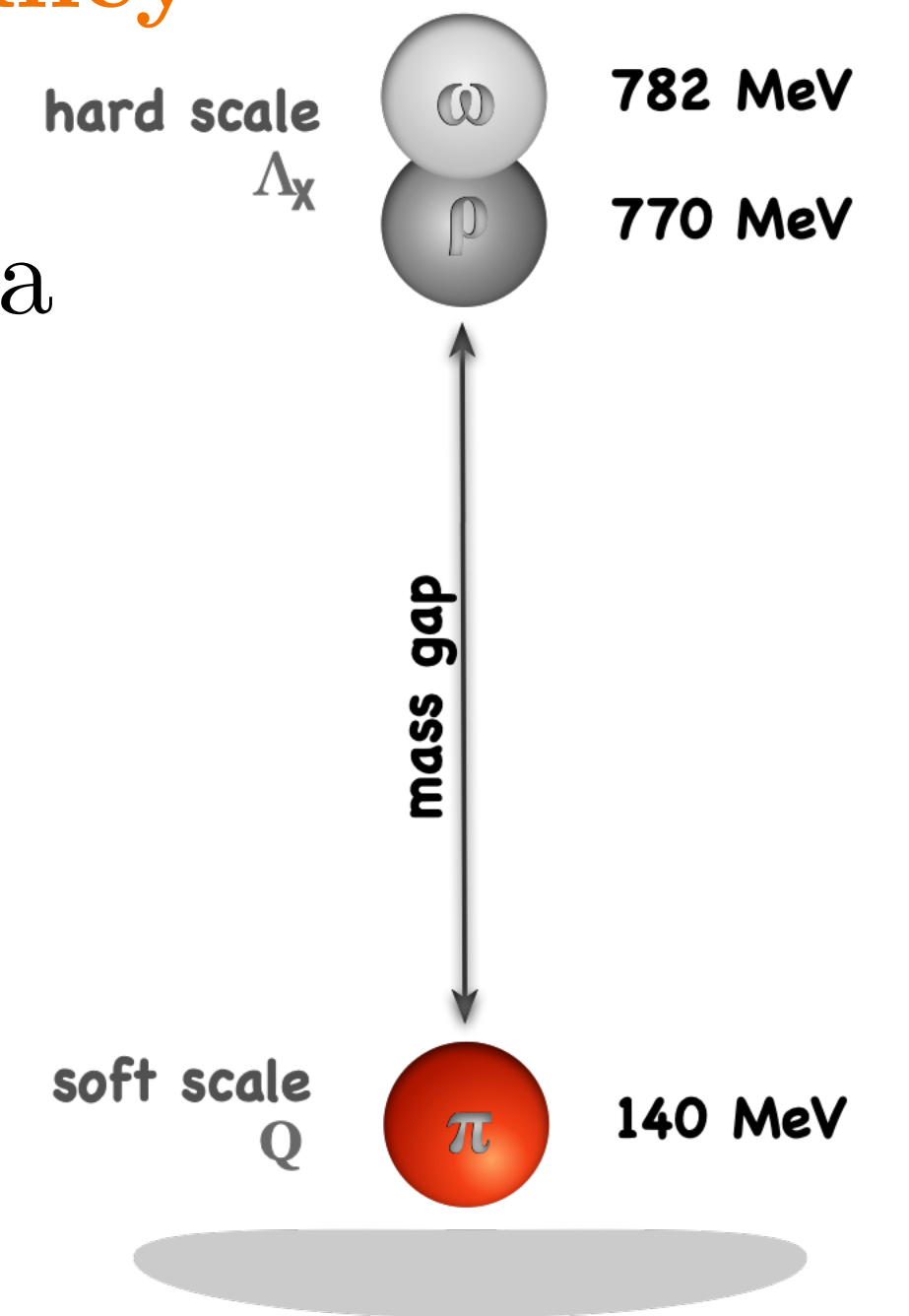
The systematicity of the *ab initio* method creates an **inferential advantage** since we claim to know something about the model discrepancy.

$$\delta M_i = y_{\text{ref}} \sum_{i=k+1}^{\infty} \gamma_i \left(\frac{Q}{\Lambda_\chi} \right)^i$$

We are uncertain about the expansion coefficients



nucleon-nucleon scattering data
(~2000 of relevance)

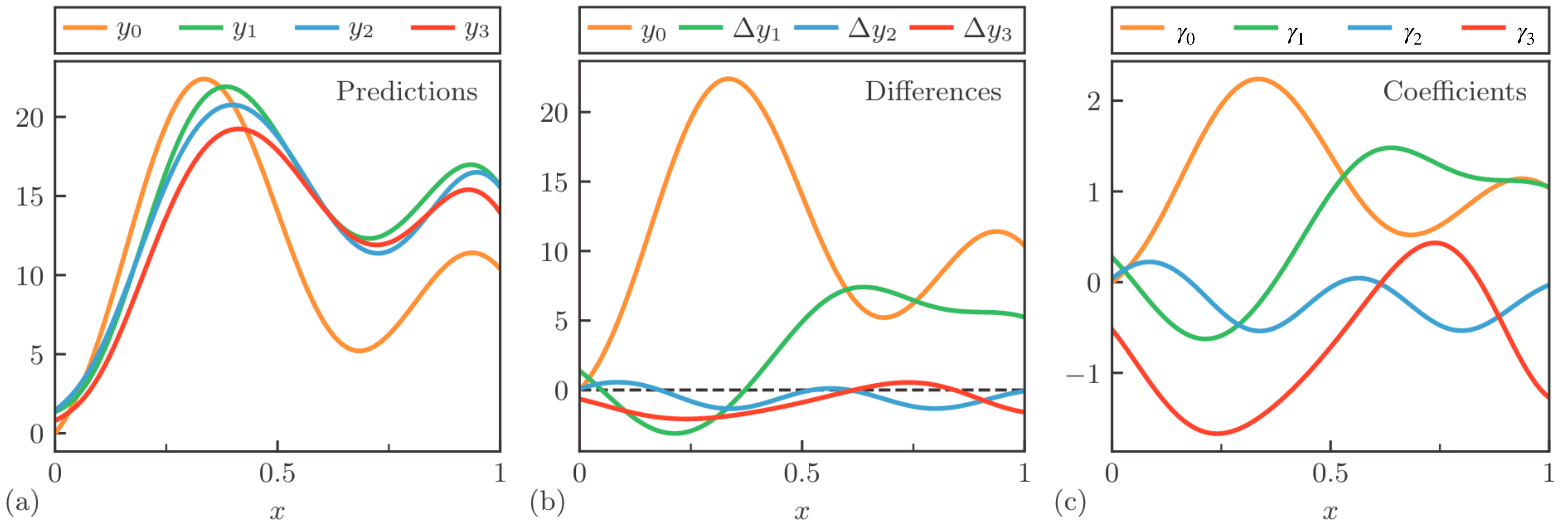


Predictions $y^{(k)}$ correlated across scattering energy & angle

adapted from J. A. Melendez, R. J. Furnstahl, D. R. Phillips, M. T. Pratola, S. Wesolowski *Phys. Rev. C* **100** (4), 044001 (2019)

QUANTIFYING CORRELATED TRUNCATION ERRORS IN ...

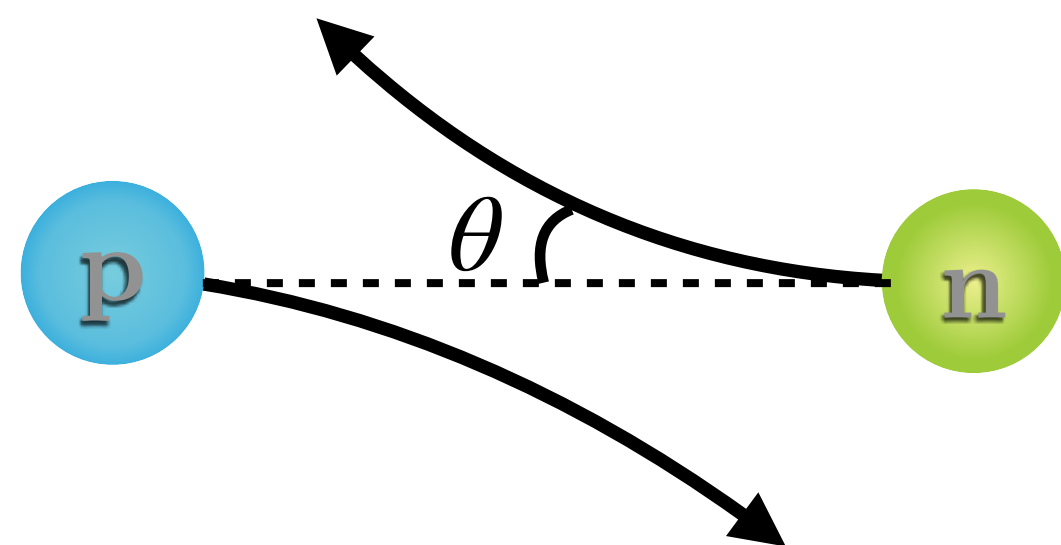
PHYSICAL REVIEW C **100**, 044001 (2019)



we use $y_{\text{ref}} = 0.5$ (Idaho-N3LO) for polarization (σ_{tot} & $\sigma(\theta)$)

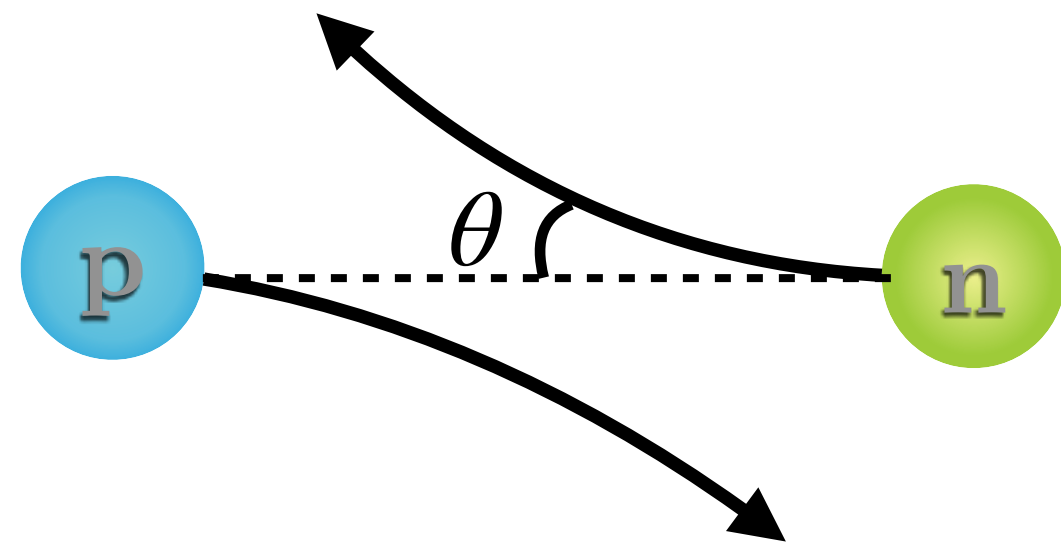
$$\gamma^{(k)}(\vec{x}_j) = \frac{y_{\text{th}}^{(k)}(\vec{\alpha}^*; \vec{x}_j) - y_{\text{th}}^{(k-1)}(\vec{\alpha}^*; \vec{x}_j)}{y_{\text{ref}}(Q/\Lambda_\chi)^k}$$

A Gaussian-process model for correlated δM



Quantify expansion coefficients by evaluating the model at consecutive chiral orders for a 2D grid of \vec{x}_j values (energies T_{lab} and angles θ).
n.b. we do this for a sensible choice of $\vec{\alpha} = \vec{\alpha}^$*

A Gaussian-process model for correlated δM

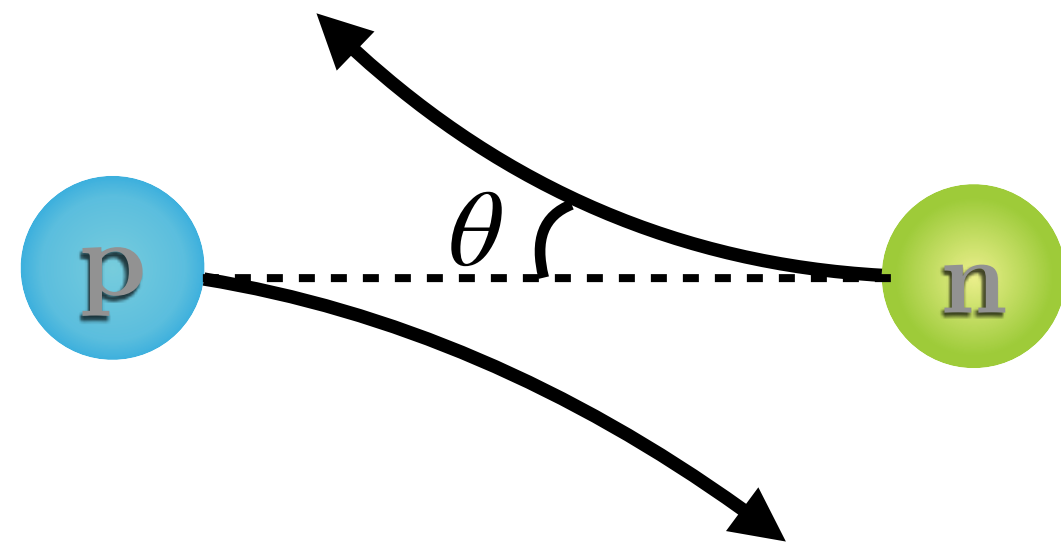


Quantify expansion coefficients by evaluating the model at consecutive chiral orders for a 2D grid of \vec{x}_j values (energies T_{lab} and angles θ).
n.b. we do this for a sensible choice of $\vec{\alpha} = \vec{\alpha}^$*

We model the correlated expansion parameters as a two-feature Gaussian process using a SQE kernel

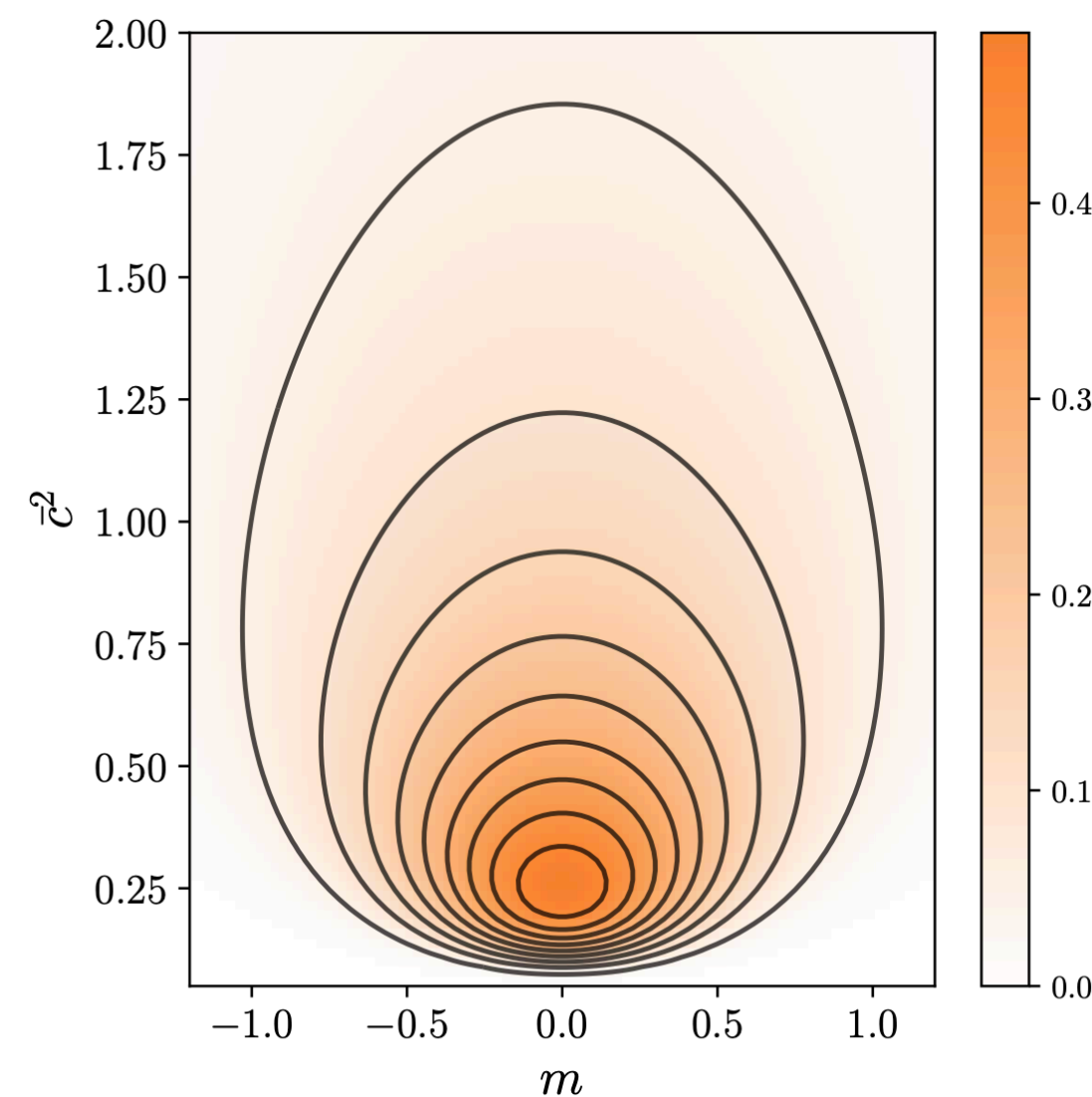
$$\gamma(\mathbf{x}) | m, \vec{\ell}, \bar{c}^2 \sim \mathcal{GP}[m, \bar{c}^2 k(\mathbf{x}', \mathbf{x}; \vec{\ell})]$$

A Gaussian-process model for correlated δM



Quantify expansion coefficients by evaluating the model at consecutive chiral orders for a 2D grid of \vec{x}_j values (energies T_{lab} and angles θ).
n.b. we do this for a sensible choice of $\vec{\alpha} = \vec{\alpha}^$*

We model the correlated expansion parameters as a two-feature Gaussian process using a SQE kernel

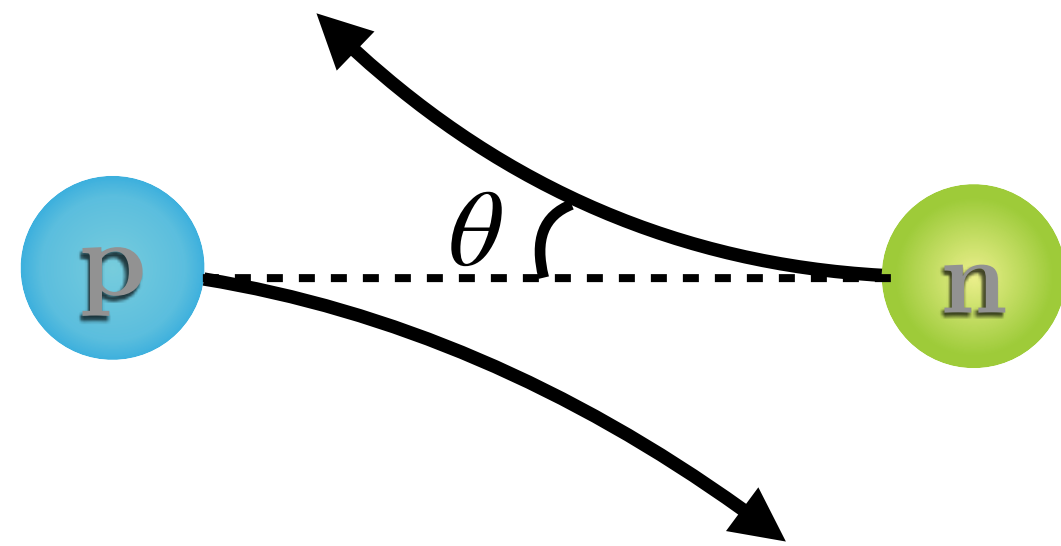


$$\gamma(\mathbf{x}) | m, \vec{\ell}, \bar{c}^2 \sim \mathcal{GP}[m, \bar{c}^2 k(\mathbf{x}', \mathbf{x}; \vec{\ell})]$$

← Our prior for the GP hyperparameters

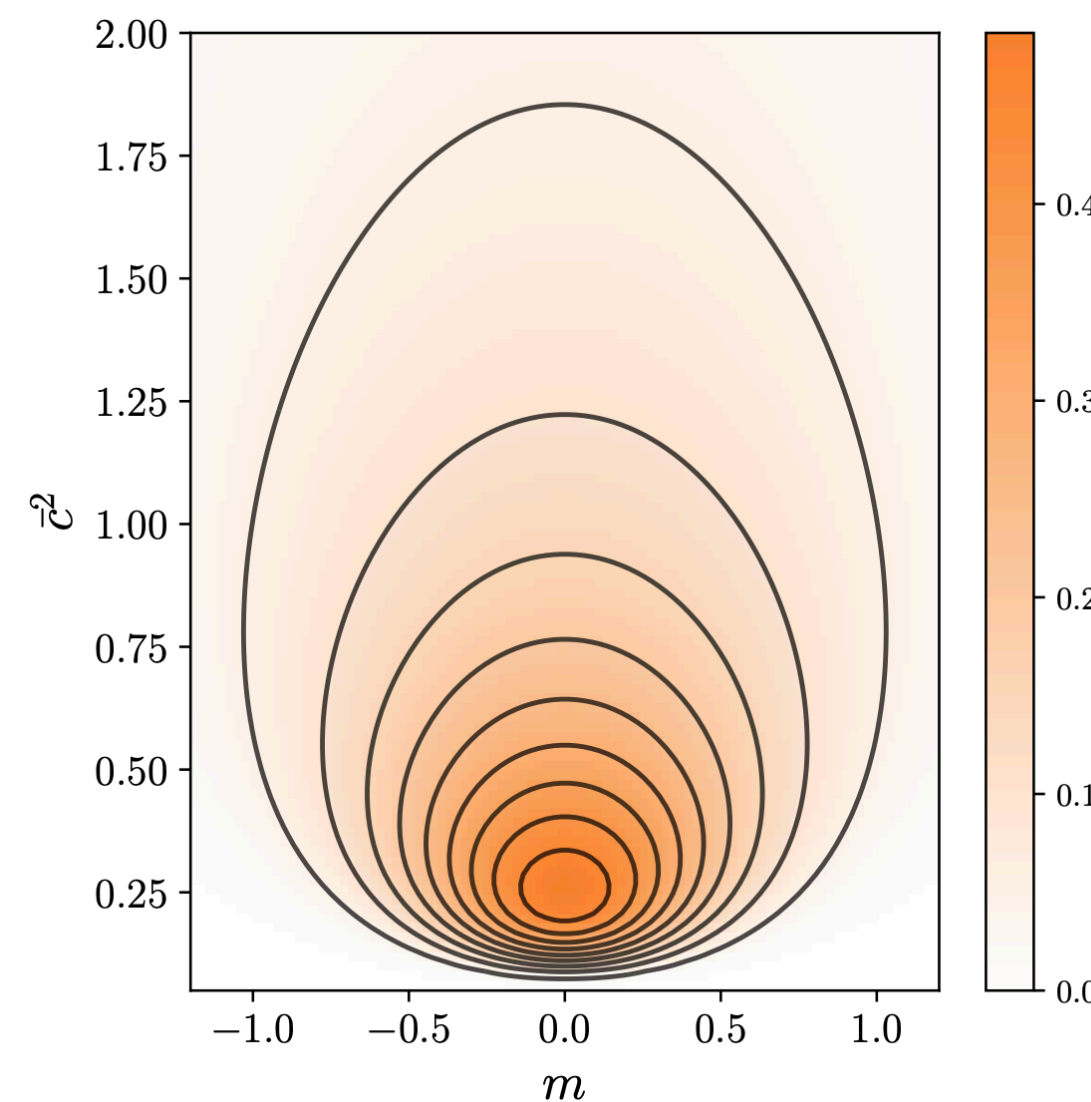
uniform distribution
for the lengthscales $\vec{\ell}$ →

A Gaussian-process model for correlated δM



Quantify expansion coefficients by evaluating the model at consecutive chiral orders for a 2D grid of \vec{x}_j values (energies T_{lab} and angles θ).
n.b. we do this for a sensible choice of $\vec{\alpha} = \vec{\alpha}^$*

We model the correlated expansion parameters as a two-feature Gaussian process using a SQE kernel

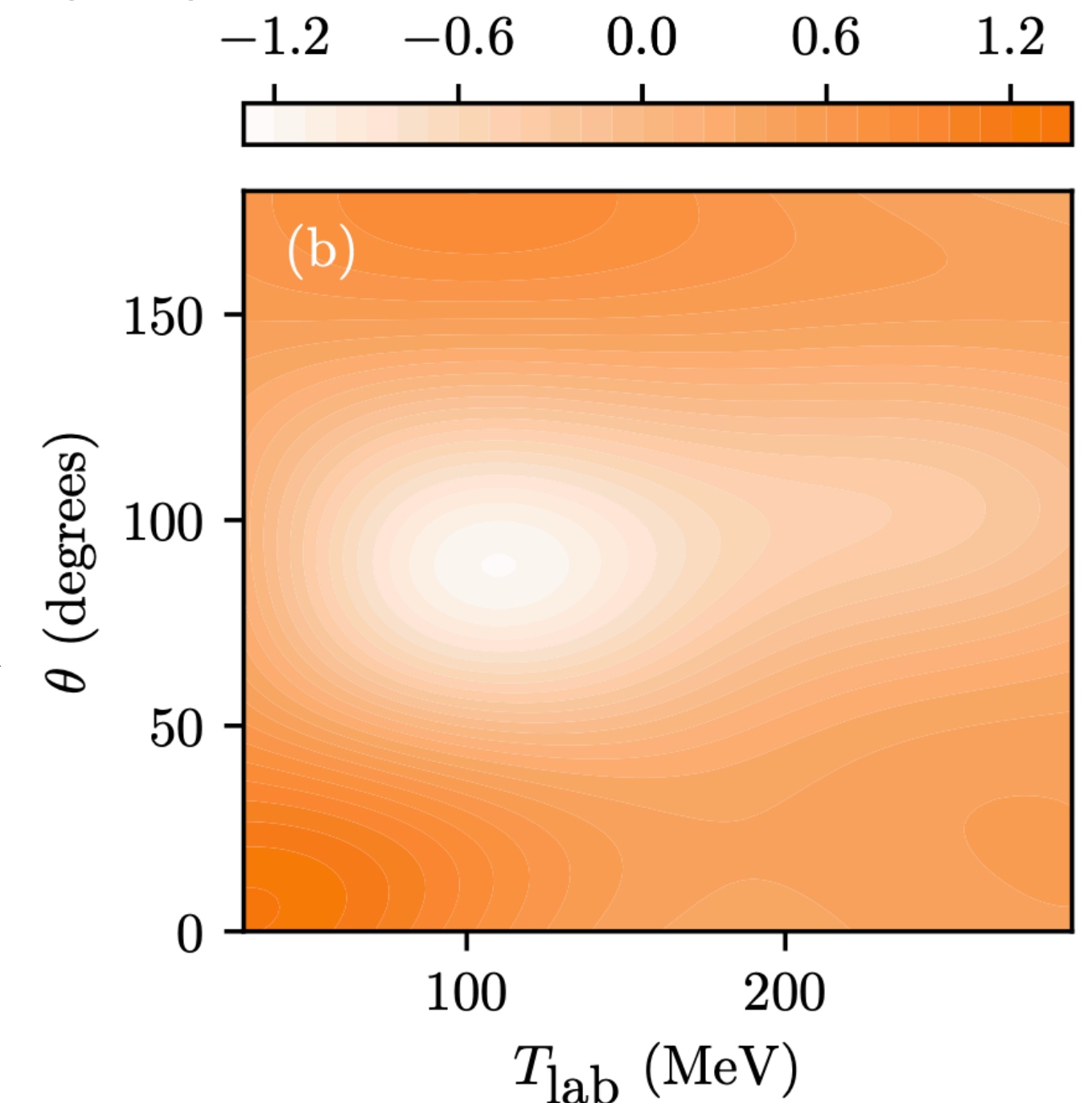


uniform distribution
for the lengthscales $\vec{\ell}$

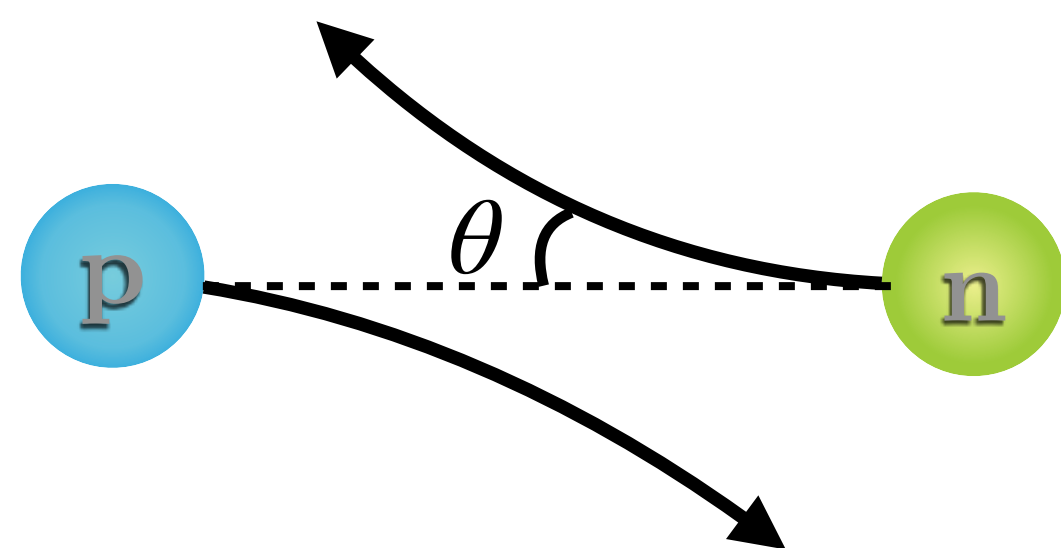
$$\gamma(\mathbf{x}) | m, \vec{\ell}, \bar{c}^2 \sim \mathcal{GP}[m, \bar{c}^2 k(\mathbf{x}', \mathbf{x}; \vec{\ell})]$$

Our prior for the GP hyperparameters

Our posterior for the EFT expansion coefficients

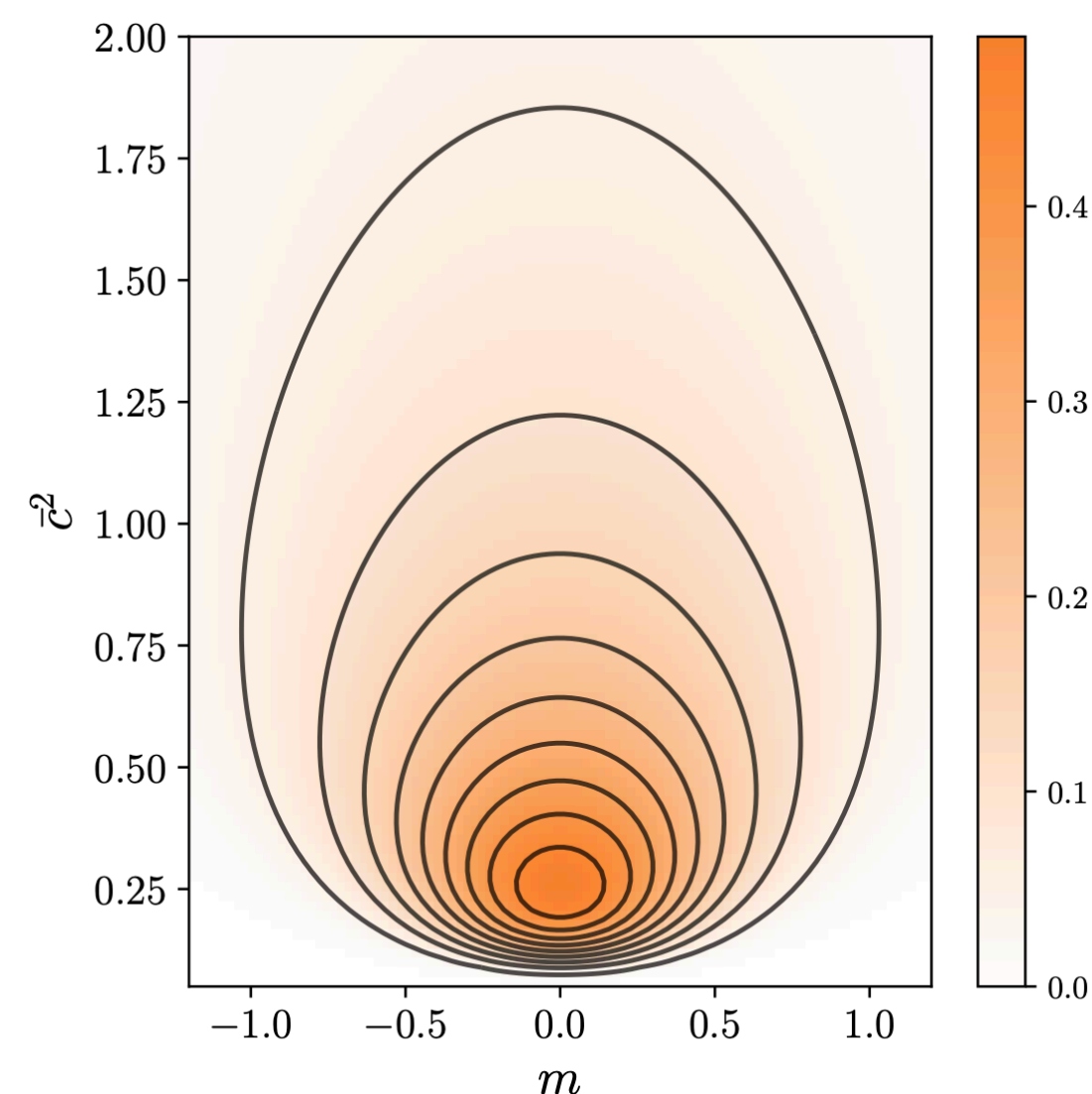


A Gaussian-process model for correlated δM



Quantify expansion coefficients by evaluating the model at consecutive chiral orders for a 2D grid of \vec{x}_j values (energies T_{lab} and angles θ).
n.b. we do this for a sensible choice of $\vec{\alpha} = \vec{\alpha}^$*

We model the correlated expansion parameters as a two-feature Gaussian process using a SQE kernel



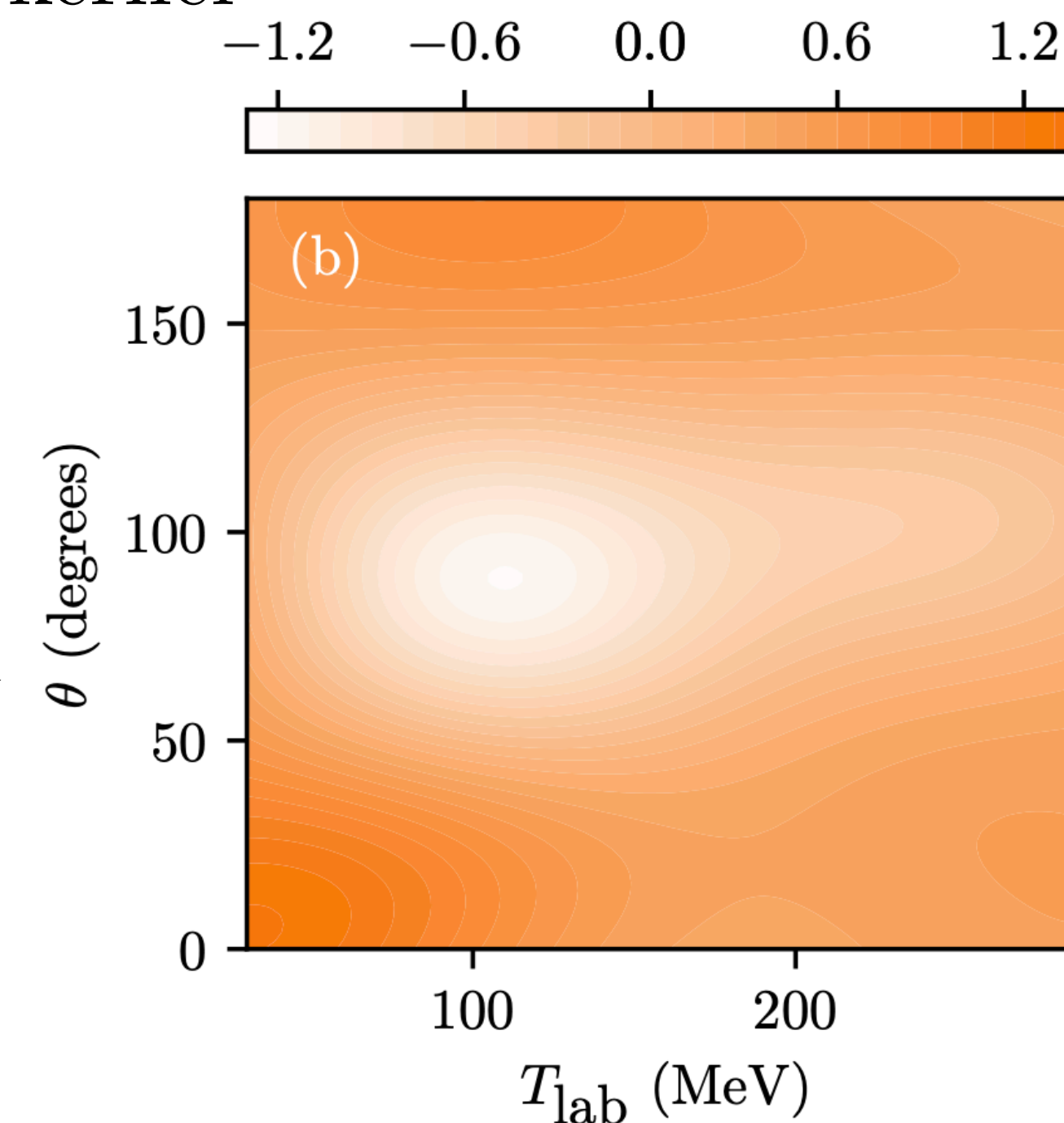
uniform distribution
for the lengthscales $\vec{\ell}$

$$\gamma(\mathbf{x}) | m, \vec{\ell}, \bar{c}^2 \sim \mathcal{GP}[m, \bar{c}^2 k(\mathbf{x}', \mathbf{x}; \vec{\ell})]$$

Our prior for the GP hyperparameters

Our posterior for the EFT expansion coefficients

$$\delta M \sim \mathcal{GP}[M, K]$$



Notation	Definition		Acronym	$N_{d,y}$	$N_{T_{\text{lab}},y}$	n_{eff}	$\widehat{\ell}_{T_{\text{lab}}} \text{ (MeV)}$	\widehat{c}^2	
σ_{tot}	total cross section		SGT	119	113	26.7	61	0.61^2	
σ_T	$\sigma_{\text{tot}}(\uparrow\downarrow) - \sigma_{\text{tot}}(\uparrow\uparrow)$		SGTT	3	3	—	—	—	
σ_L	$\sigma_{\text{tot}}(\leftarrow\rightarrow) - \sigma_{\text{tot}}(\rightarrow\leftarrow)$		SGTL	4	4	3.7	59	2.28^2	
Notation	Tensor	Illustration	Acronym	$N_{d,y}$	$N_{T_{\text{lab}},y}$	n_{eff}	$\widehat{\ell}_{T_{\text{lab}}} \text{ (MeV)}$	$\widehat{\ell}_\theta \text{ (deg)}$	\widehat{c}^2
$\sigma(\theta)$	I_{0000}		DSG	1207	68	383.8	73	39	0.56^2
$A(\theta)$	D_{s0k0}		A	5	1	5.0	70	37	0.66^2
$A_t(\theta)$	K_{0ks0}		AT	30	2	23.0	60	36	0.61^2
$A_{yy}(\theta)$	A_{00nn}		AYY	58	4	19.4	60	33	0.97^2
$A_{zz}(\theta)$	A_{00kk}		AZZ	45	2	12.1	120	37	1.45^2
$D(\theta)$	D_{n0n0}		D	13	1	4.7	68	26	0.6^2
$D_t(\theta)$	K_{0nn0}		DT	36	5	33.6	41	45	0.52^2
$D_x^z(\theta)$	D_{0s0k}		DOSK	8	1	3.1	59	29	0.52^2
$A_y(\theta)$	P_{n000}		P	503	28	269.6	65	32	0.35^2
$N_{zz}^y(\theta)$	N_{0nkk}		NNKK	8	1	8.0	45	25	0.22^2
$N_{zy}^x(\theta)$	N_{0skn}		NSKN	12	1	11.0	83	36	0.85^2
$N_{xy}^x(\theta)$	N_{0ssn}		NSSN	4	1	4.0	78	31	0.62^2
$R(\theta)$	D_{s0s0}		R	5	1	5.0	74	32	0.67^2
$R_t(\theta)$	K_{0ss0}		RT	29	3	24.2	87	28	0.59^2
$R'_t(\theta)$	K_{0sk0}		RPT	1	1	1.0	62	35	0.54^2



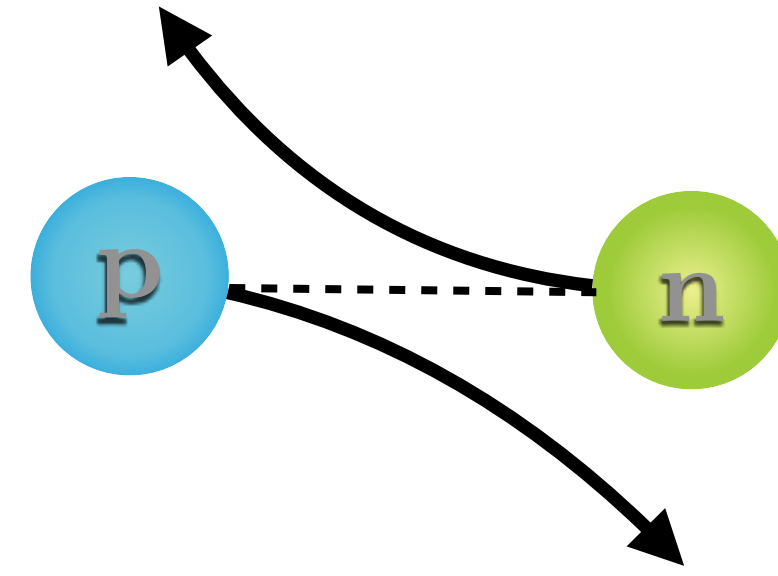
We construct a model discrepancy for each observable type (y) we condition our inference on.

correlation lengths for np scattering energies and angles in the ranges **40–120 MeV** and **25–45 degrees**.

small marginal variances

The likelihood

$$\begin{aligned} \text{pr}(\mathcal{D}|\vec{\alpha}, I) &= \prod_{y=1}^{N_y} \text{pr}(\mathcal{D}_y|\vec{\alpha}, I_y), \\ \text{pr}(\mathcal{D}_y|\vec{\alpha}, I_y) &\propto \exp\left[-\frac{\mathbf{r}_y^T(\vec{\alpha})(\boldsymbol{\Sigma}_{\text{exp},y} + \boldsymbol{\Sigma}_{\text{th},y})^{-1}\mathbf{r}_y(\vec{\alpha})}{2}\right] \\ r_{y,j}(\vec{\alpha}) &= [y_{\text{exp}}(\vec{x}_j) - y_{\text{th}}^{(k)}(\vec{\alpha}; \vec{x}_j)] \\ \boldsymbol{\Sigma}_{\text{th}} &= \begin{bmatrix} \boldsymbol{\Sigma}_{\text{th},1} & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Sigma}_{\text{th},2} & \cdots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \cdots & \boldsymbol{\Sigma}_{\text{th},N_y} \end{bmatrix} \\ (\boldsymbol{\Sigma}_{\text{th},y})_{mn} &= \text{cov}[\delta y_{\text{th}}^{(k)}(\vec{x}_m), \delta y_{\text{th}}^{(k)}(\vec{x}_n)] \end{aligned}$$



The likelihood

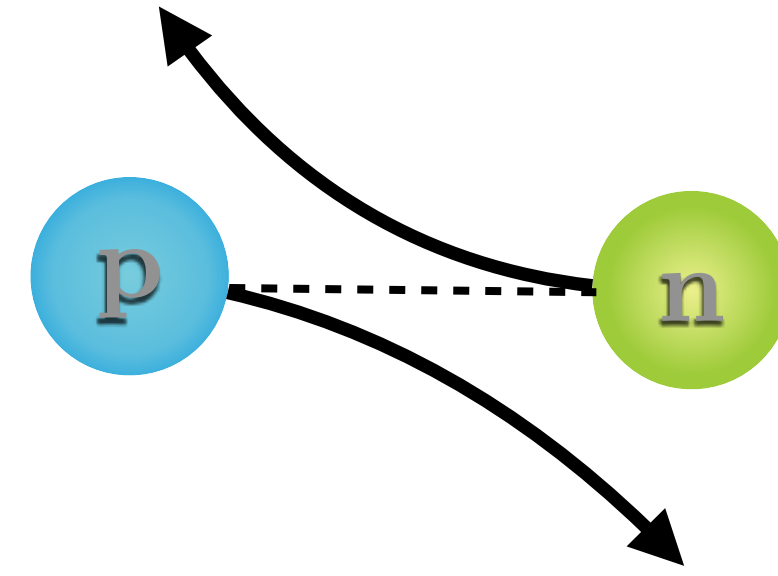
$$\text{pr}(\mathcal{D}|\vec{\alpha}, I) = \prod_{y=1}^{N_y} \text{pr}(\mathcal{D}_y|\vec{\alpha}, I_y),$$

$$\text{pr}(\mathcal{D}_y|\vec{\alpha}, I_y) \propto \exp \left[-\frac{\mathbf{r}_y^T(\vec{\alpha})(\boldsymbol{\Sigma}_{\text{exp},y} + \boldsymbol{\Sigma}_{\text{th},y})^{-1}\mathbf{r}_y(\vec{\alpha})}{2} \right]$$

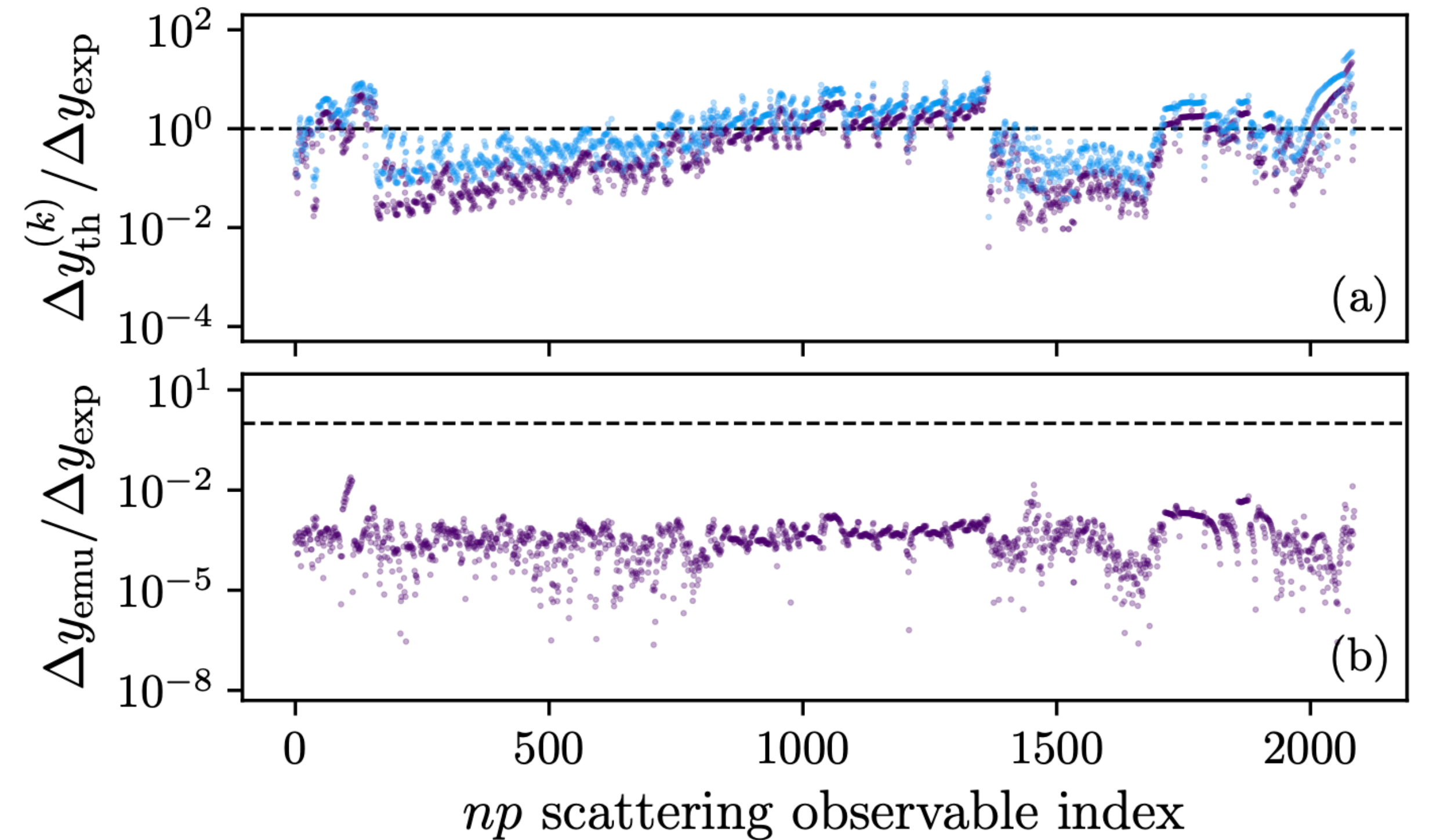
$$r_{y,j}(\vec{\alpha}) = [y_{\text{exp}}(\vec{x}_j) - y_{\text{th}}^{(k)}(\vec{\alpha}; \vec{x}_j)]$$

$$\boldsymbol{\Sigma}_{\text{th}} = \begin{bmatrix} \boldsymbol{\Sigma}_{\text{th},1} & \mathbf{0} & \cdots & \mathbf{0} \\ \mathbf{0} & \boldsymbol{\Sigma}_{\text{th},2} & \cdots & \mathbf{0} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{0} & \mathbf{0} & \cdots & \boldsymbol{\Sigma}_{\text{th},N_y} \end{bmatrix}$$

$$(\boldsymbol{\Sigma}_{\text{th},y})_{mn} = \text{cov}[\delta y_{\text{th}}^{(k)}(\vec{x}_m), \delta y_{\text{th}}^{(k)}(\vec{x}_n)]$$



We emulate all ~ 2000 cross sections in \mathcal{D} .
 Takes 1s (non-threaded + Google JAX)
 to evaluate likelihood and its derivatives.

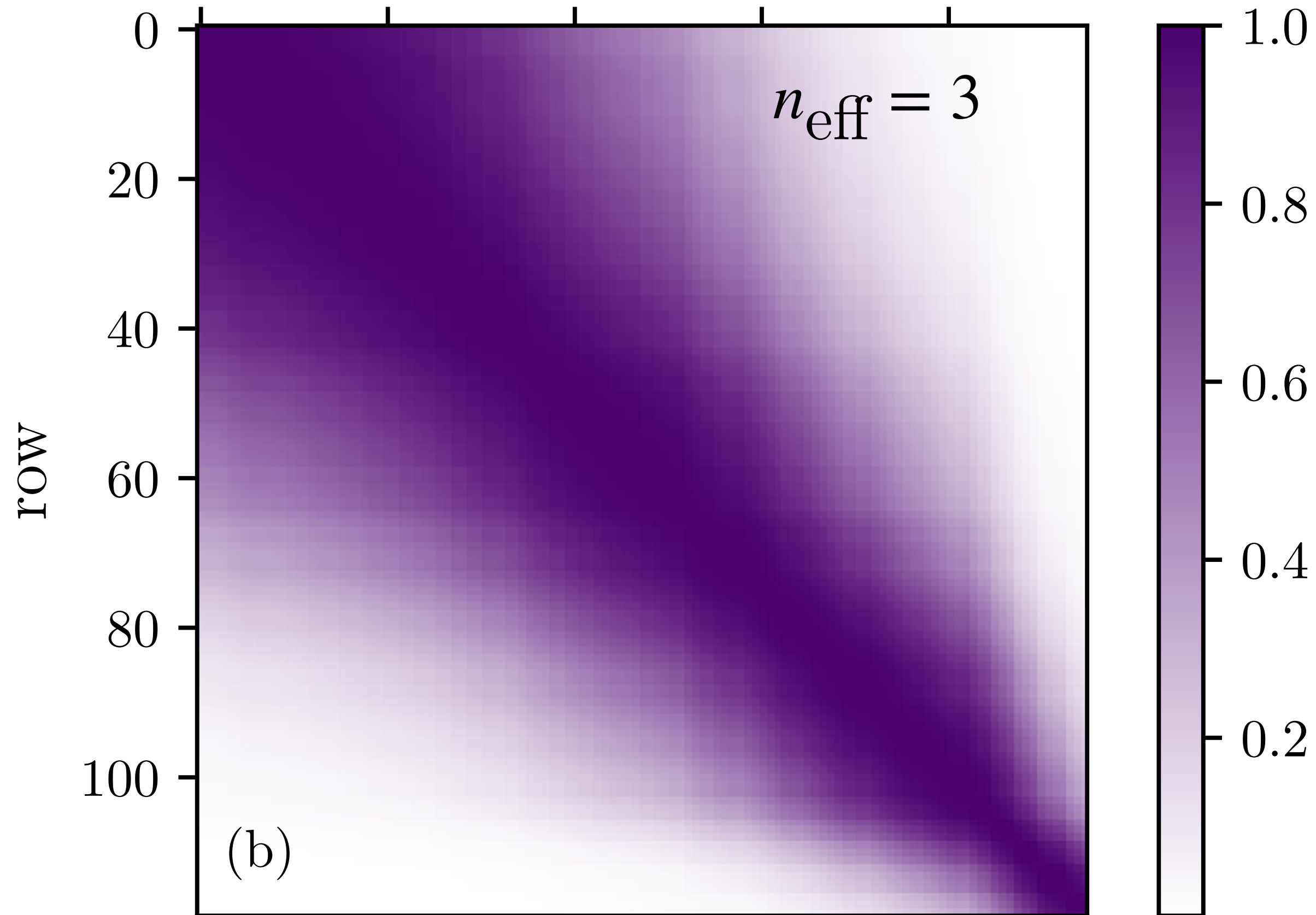


Effects of a small marginal variance in the model discrepancy

total cross section ($N_{\text{data}} = 119$)

column

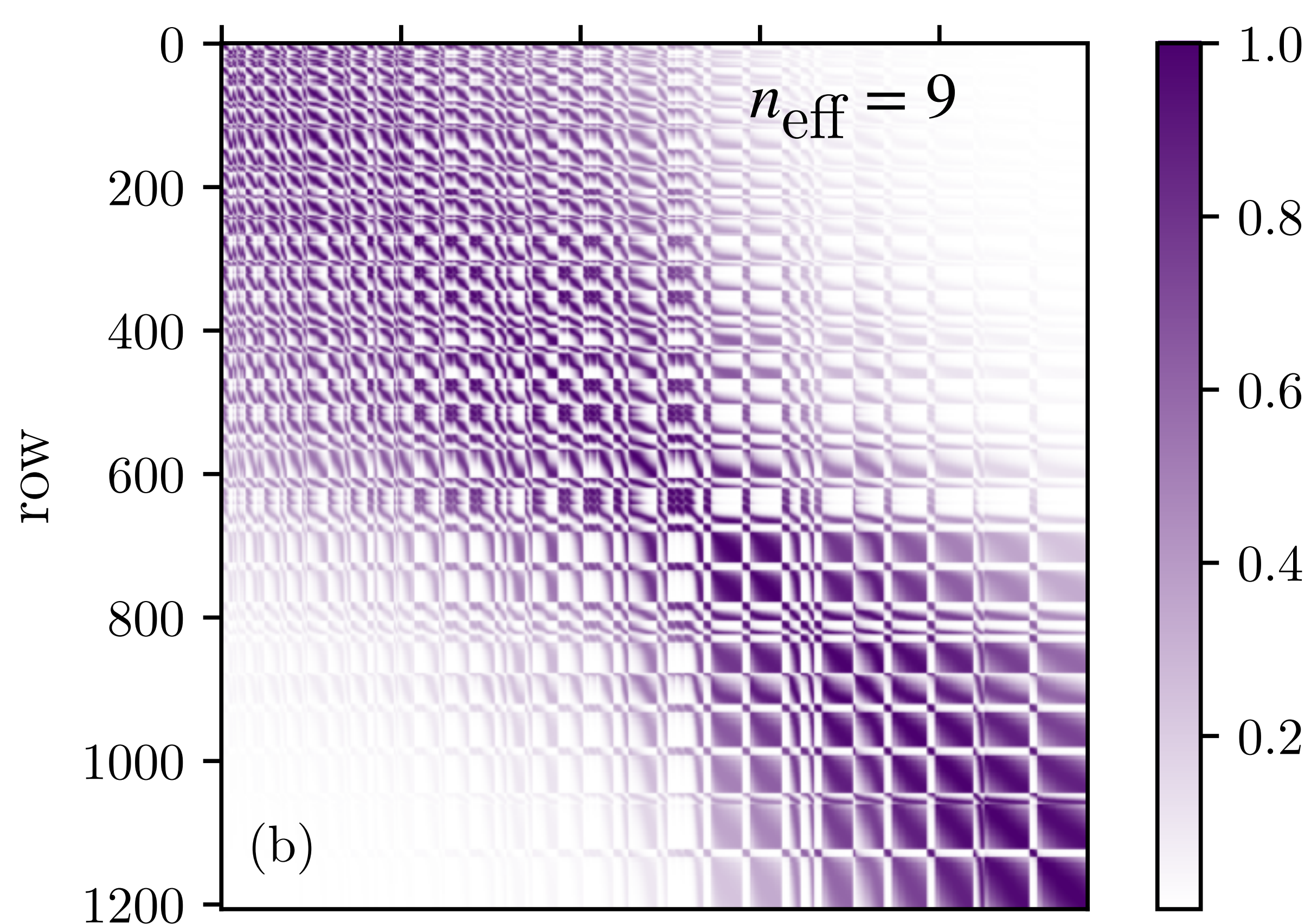
0 25 50 75 100



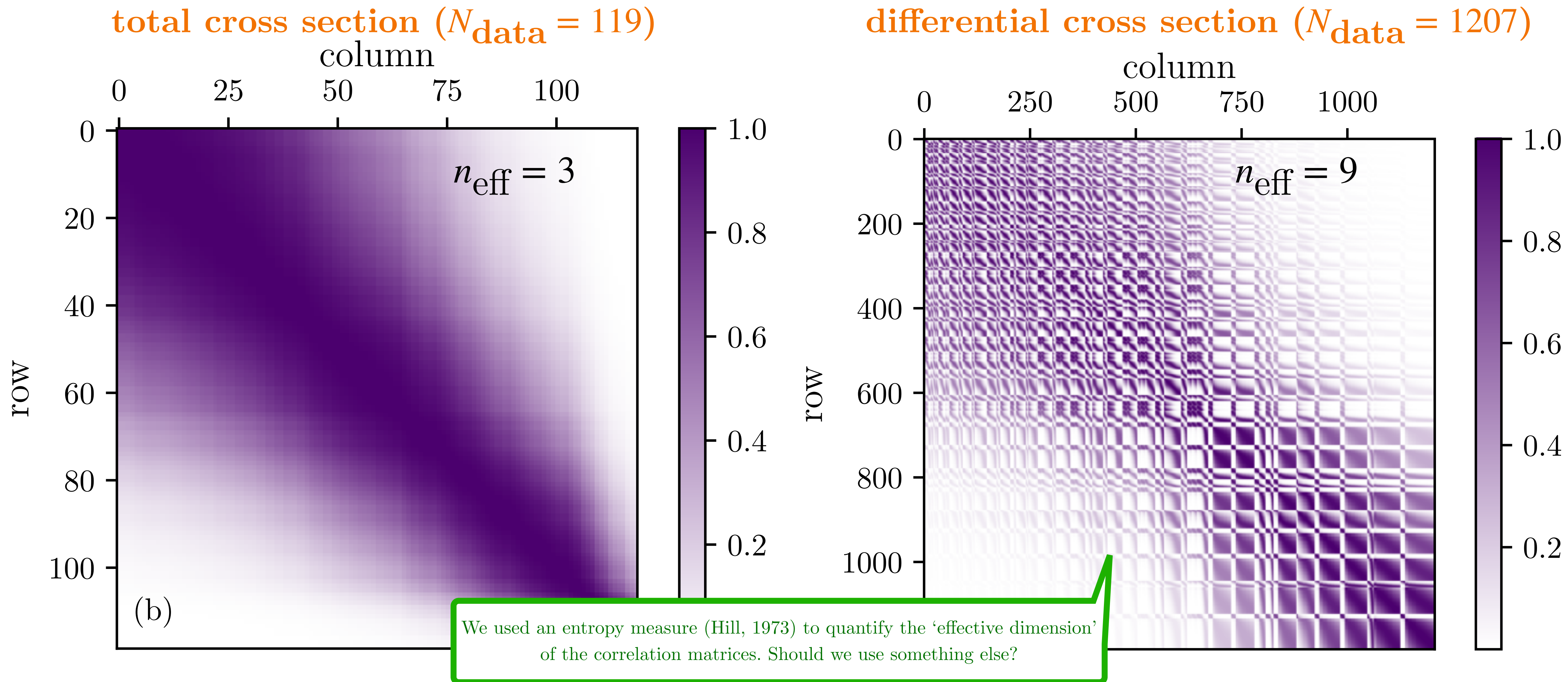
differential cross section ($N_{\text{data}} = 1207$)

column

0 250 500 750 1000



Effects of a small marginal variance in the model discrepancy

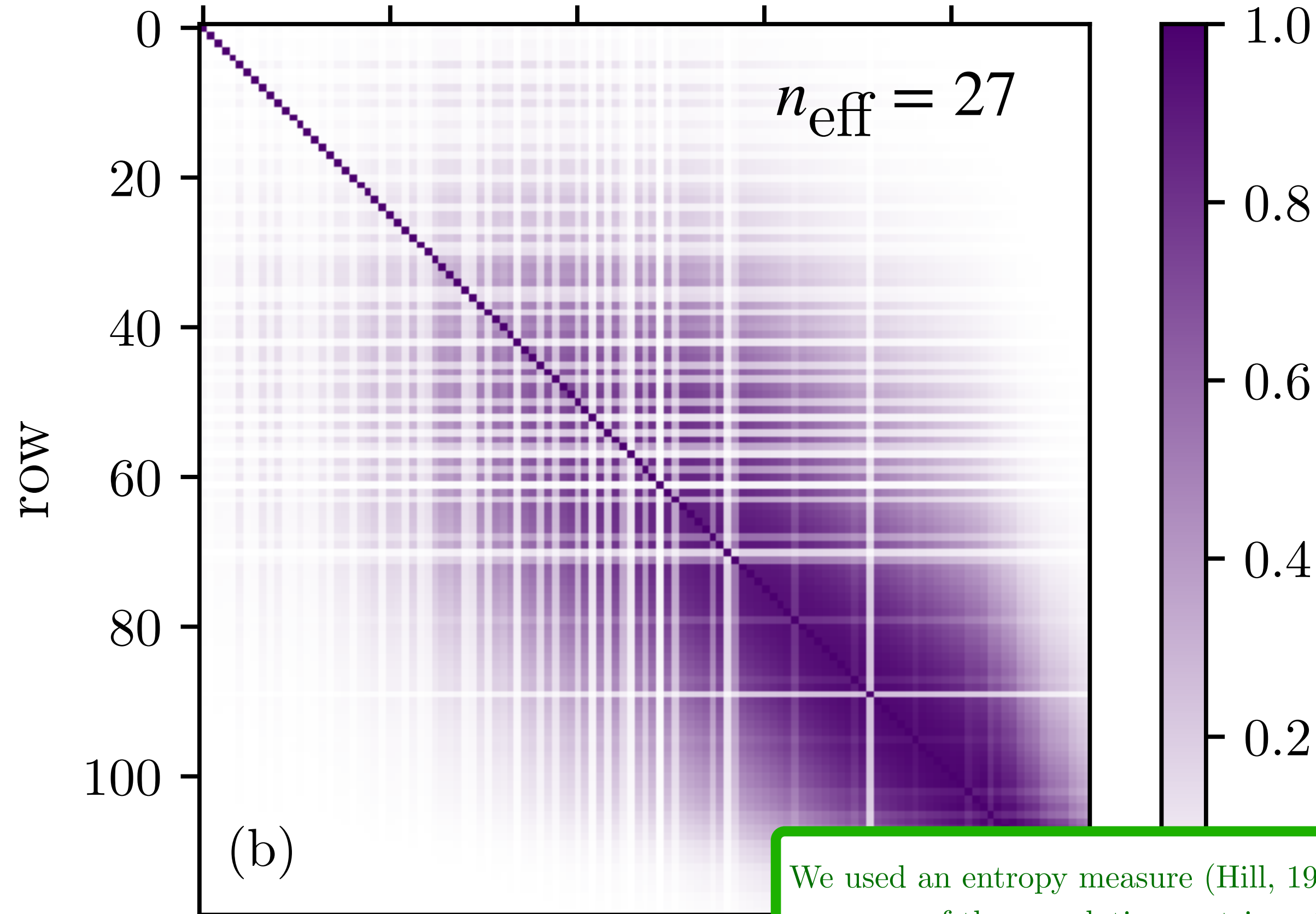


Effects of a small marginal variance in the model discrepancy

total cross section ($N_{\text{data}} = 119$)

column

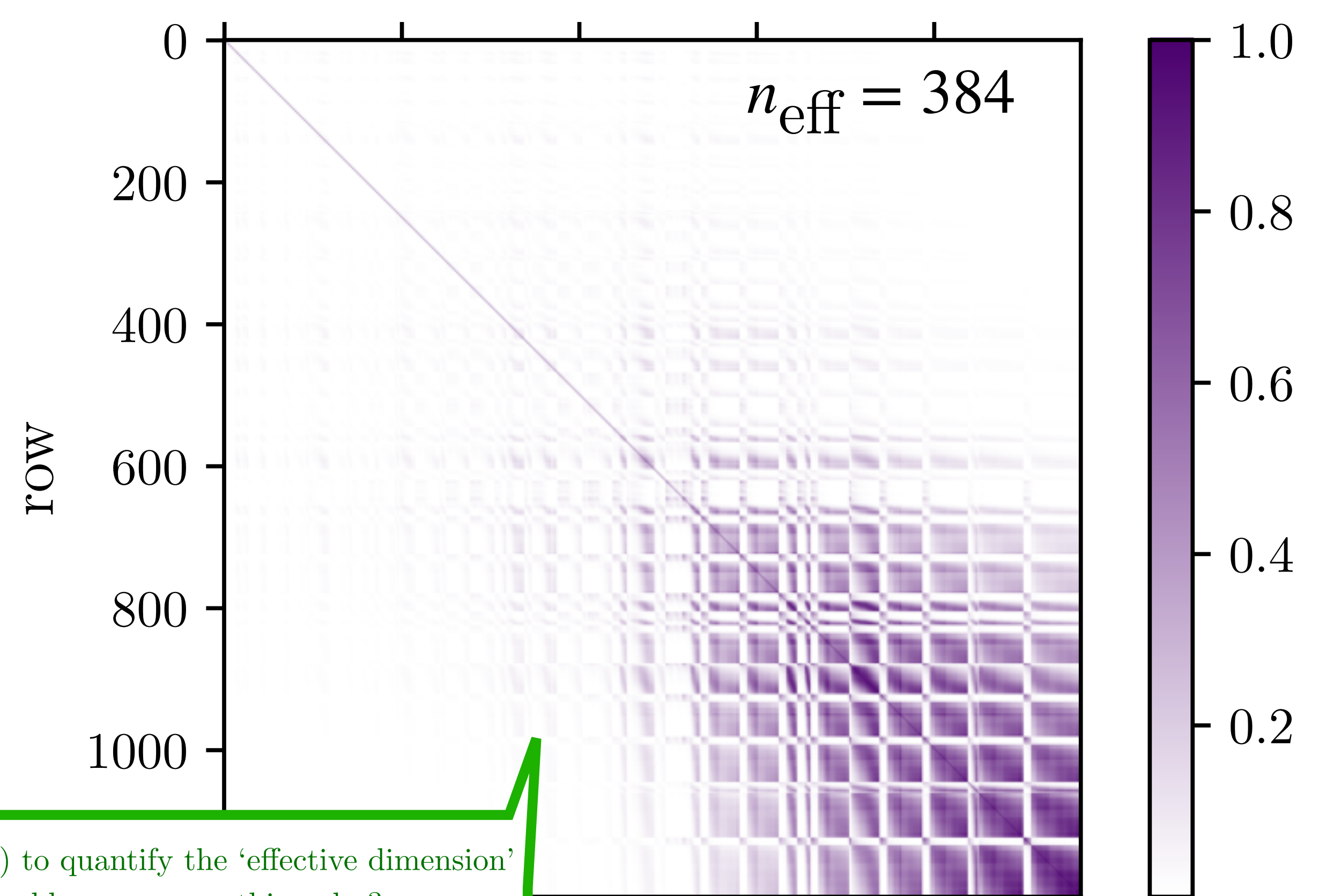
0 25 50 75 100



differential cross section ($N_{\text{data}} = 1207$)

column

0 250 500 750 1000



We used an entropy measure (Hill, 1973) to quantify the 'effective dimension' of the correlation matrices. Should we use something else?

Bayes' rule: from likelihood & prior to posterior

- Collect N data points that we gather in a data vector D
- To explain the data, propose some model M , depending on parameters $\vec{\alpha}$
- Apply Bayes' rule

$$p(\vec{\alpha} | D, M, I) = \frac{\text{Posterior} \quad \text{Likelihood} \quad \text{Prior}}{\text{Marginal likelihood}} = \frac{p(D | \vec{\alpha}, M, I) \cdot p(\vec{\alpha} | M, I)}{p(D | M, I)}$$

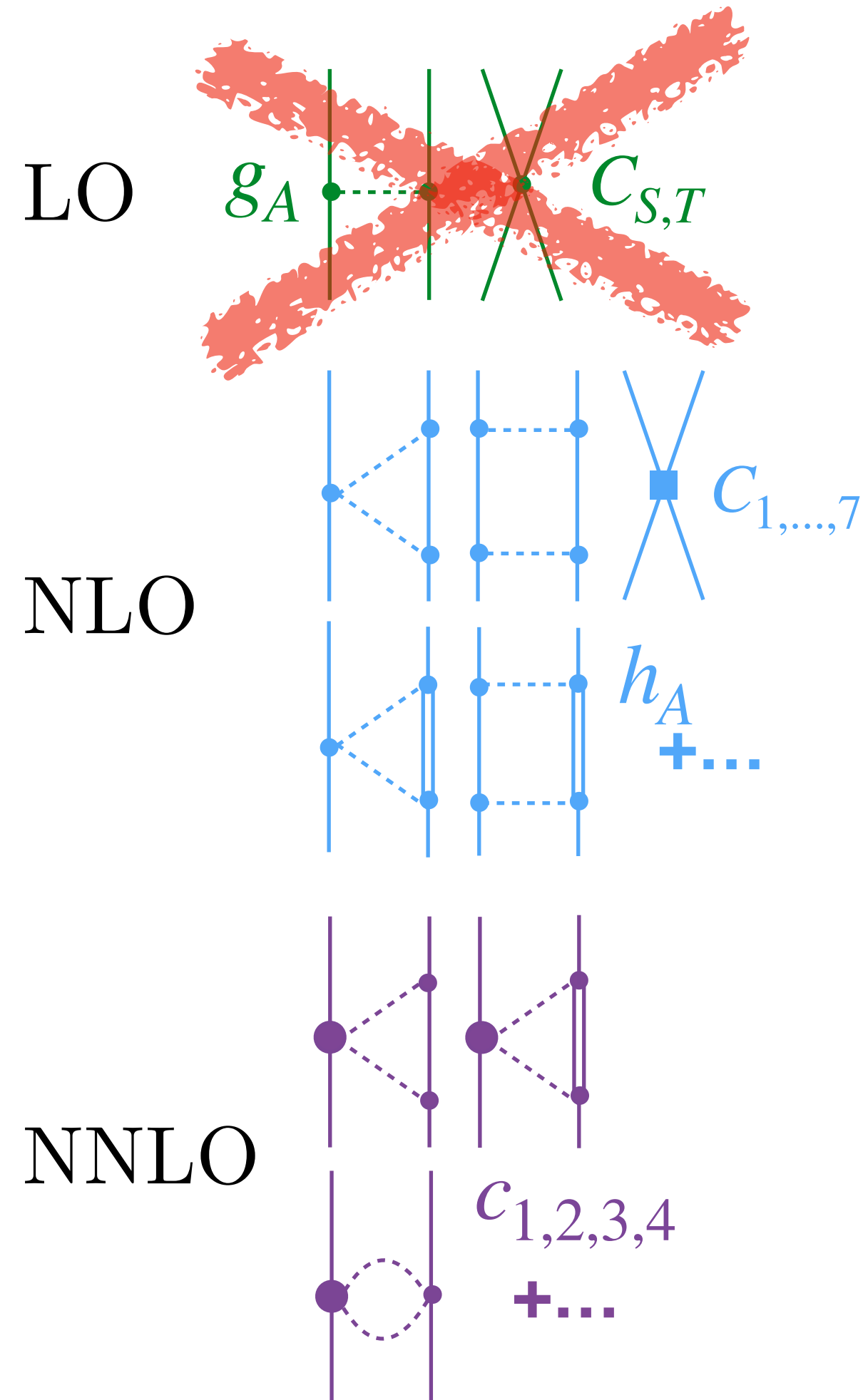


most likely not Rev. T. Bayes

- The **prior** encodes our knowledge about the parameter values before analyzing the data
- The **likelihood** is the probability of the data given a set of parameters
- The **marginal likelihood** (or model evidence) provides normalization of the posterior
- The **posterior** is the complete inference and resulting probability density for the parameters $\vec{\alpha}$

Setting up the prior for the model params'

$$p(\vec{\alpha} | I) = p(\vec{\alpha}_{NN} | I) \cdot p(\vec{\alpha}_{\pi N} | I)$$

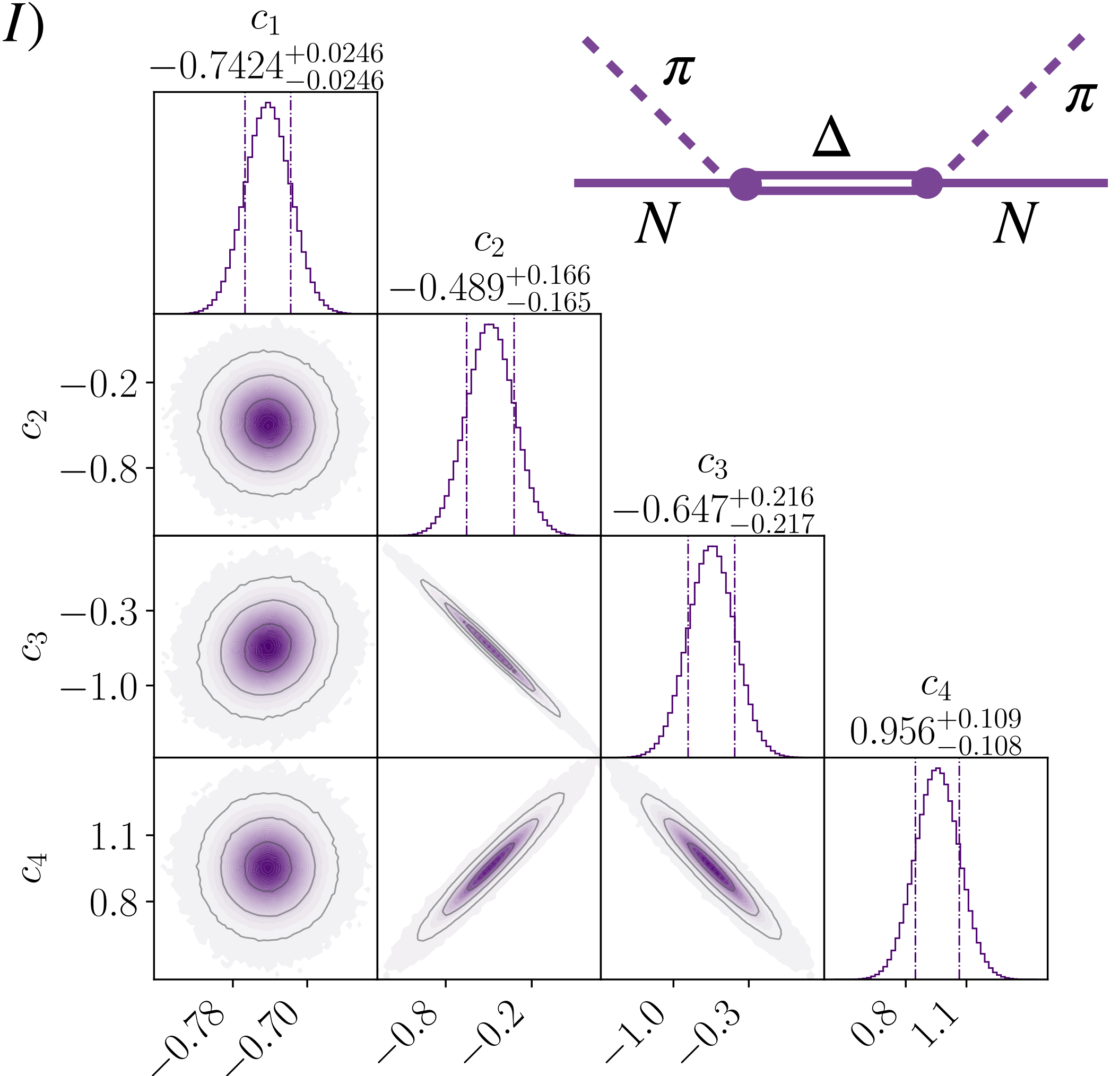


NN contacts: iid normal

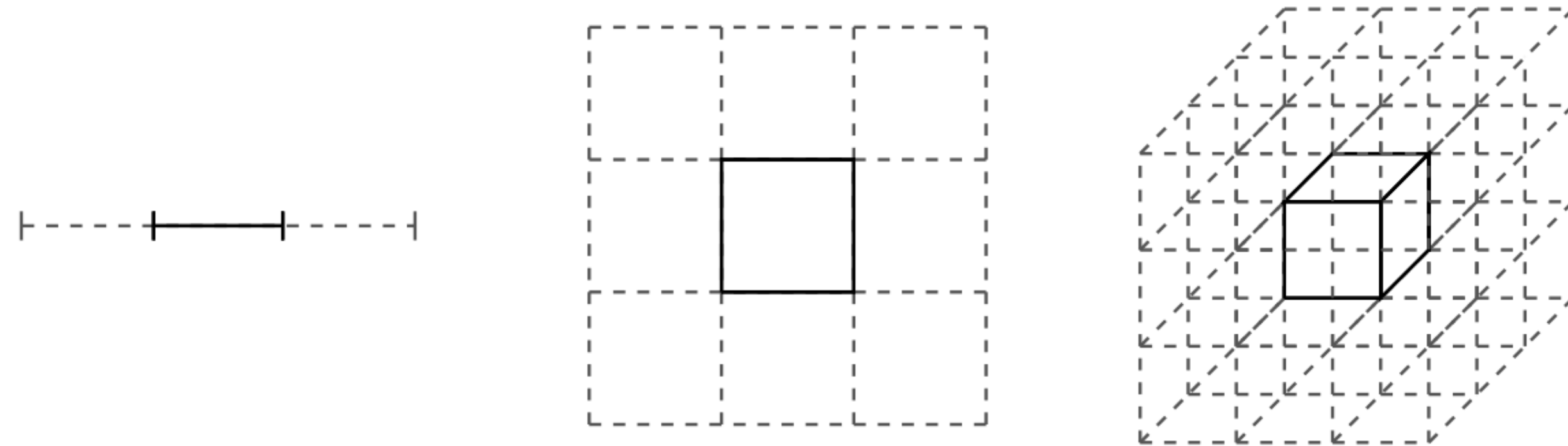
$$p(\vec{\alpha} | I) = p(\vec{\alpha}_{NN} | I) = \mathcal{N}(\mathbf{0}, \mathbf{1} \cdot 5^2)$$

πN exchange: Roy-Steiner

$$p(\vec{\alpha} | I) = p(\vec{\alpha}_{\pi N} | I)$$



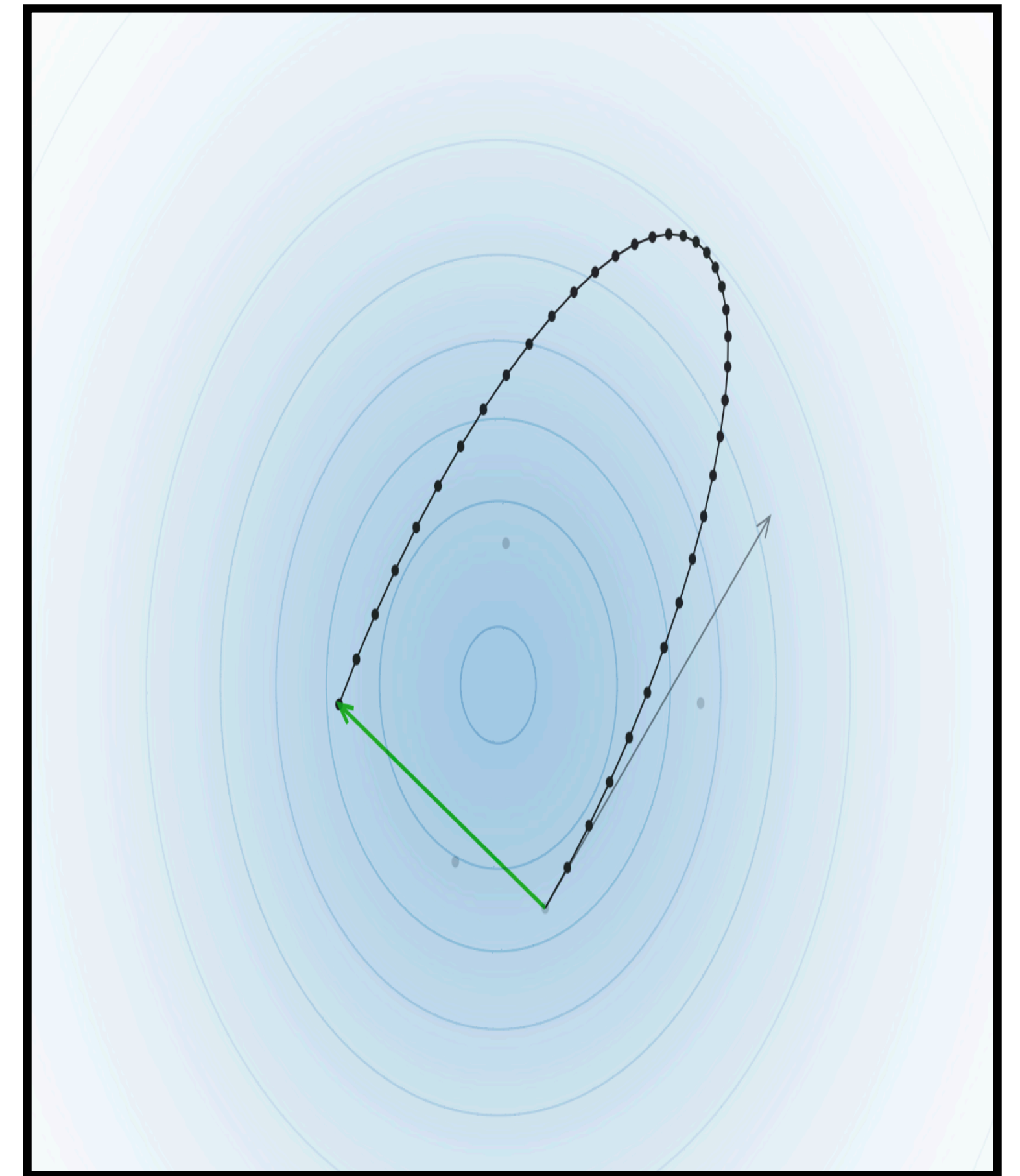
Hamiltonian Monte Carlo



credit: M. Betancourt arXiv:1701.02434

The LEC posterior is often multivariate (30 np/pp LECs at N3LO). Naive “guess and check” (random walk metropolis) will fail exponentially. We use Hamiltonian Monte Carlo (HMC) to take long jumps in parameter space while staying in regions with high probability mass.

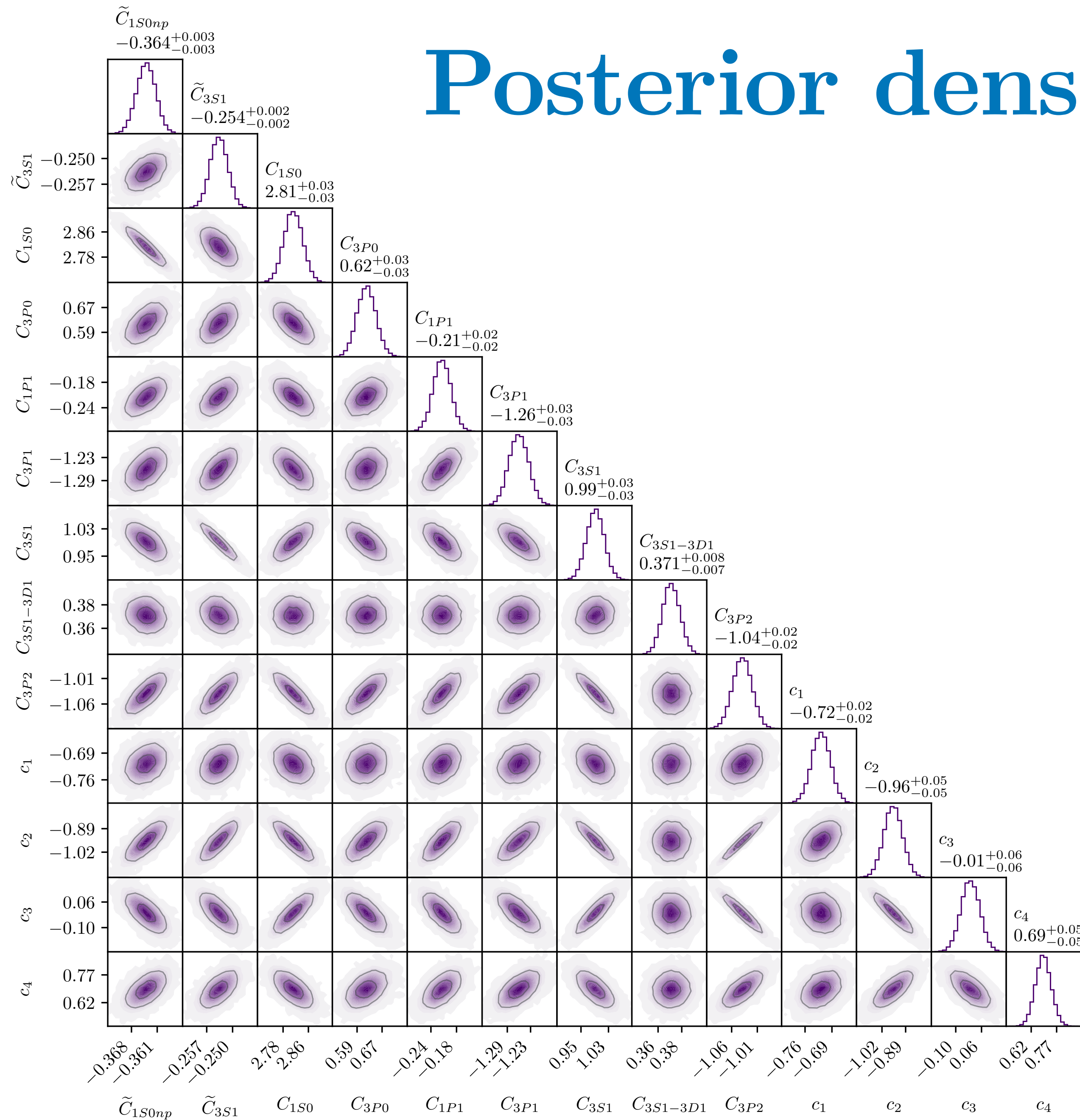
See **D. Mondal’s talk for more on HMC**



credit: <https://chi-feng.github.io/mcmc-demo/>

S. Duane, et al. Phys. Lett B **195**, 216 (1986)
 I. Svensson, et al Phys. Rev. C **105**, 014004(2022)
 I. Svensson, et al Phys. Rev. C **107**, 014001(2023)

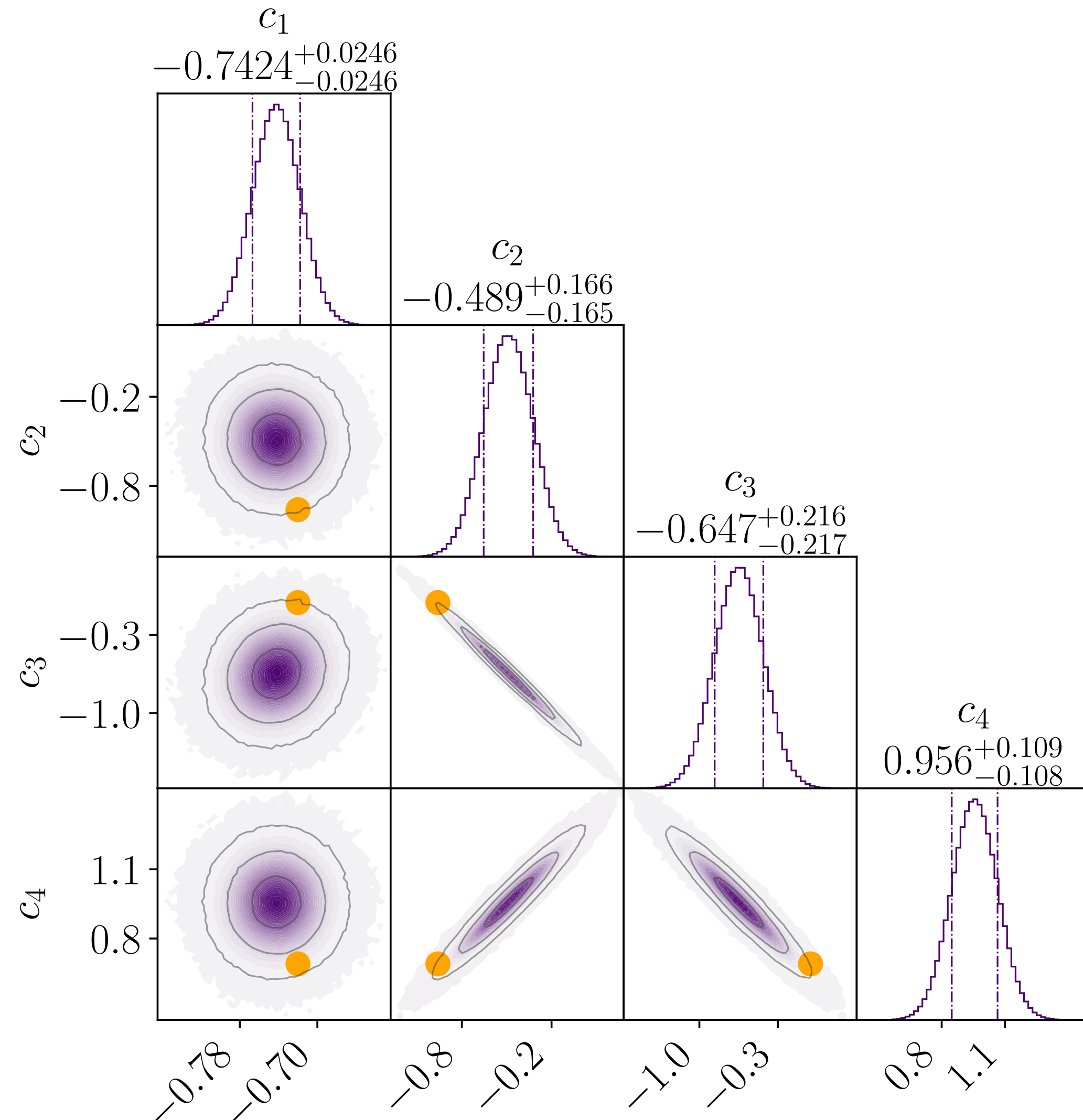
Posterior density (correlated δM)



credible ranges for the model parameters increase by a factor 1.5 when accounting for correlations in the model discrepancy

np data reduces the prior credible ranges for the πN parameters c_1, c_2, c_3, c_4 by a factor 3-5

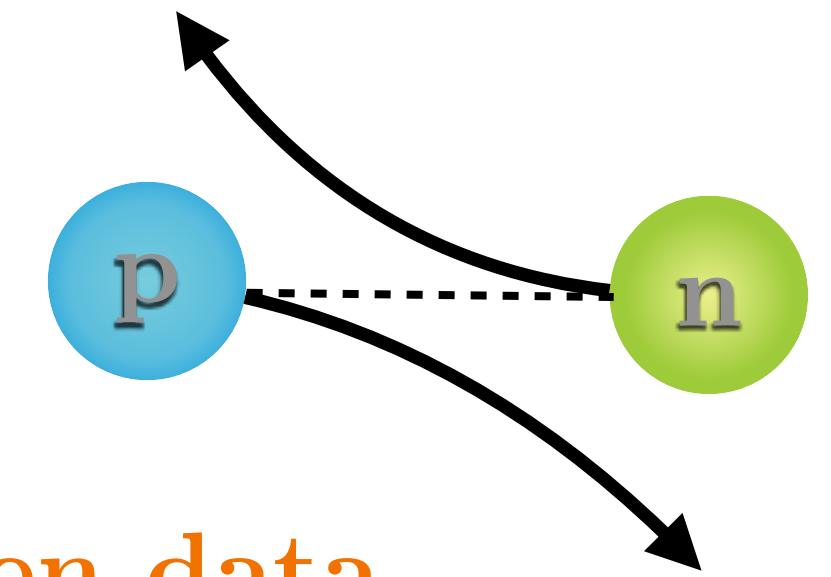
The πN prior vs posterior MAP



The Mahalanobis distance ($D_{MD}^2 = 9.95$) between the πN posterior MAP and prior (normal) mean is far enough to be outside 95% of the prior probability mass.

We need a better understanding of the model discrepancy and the underlying EFT.

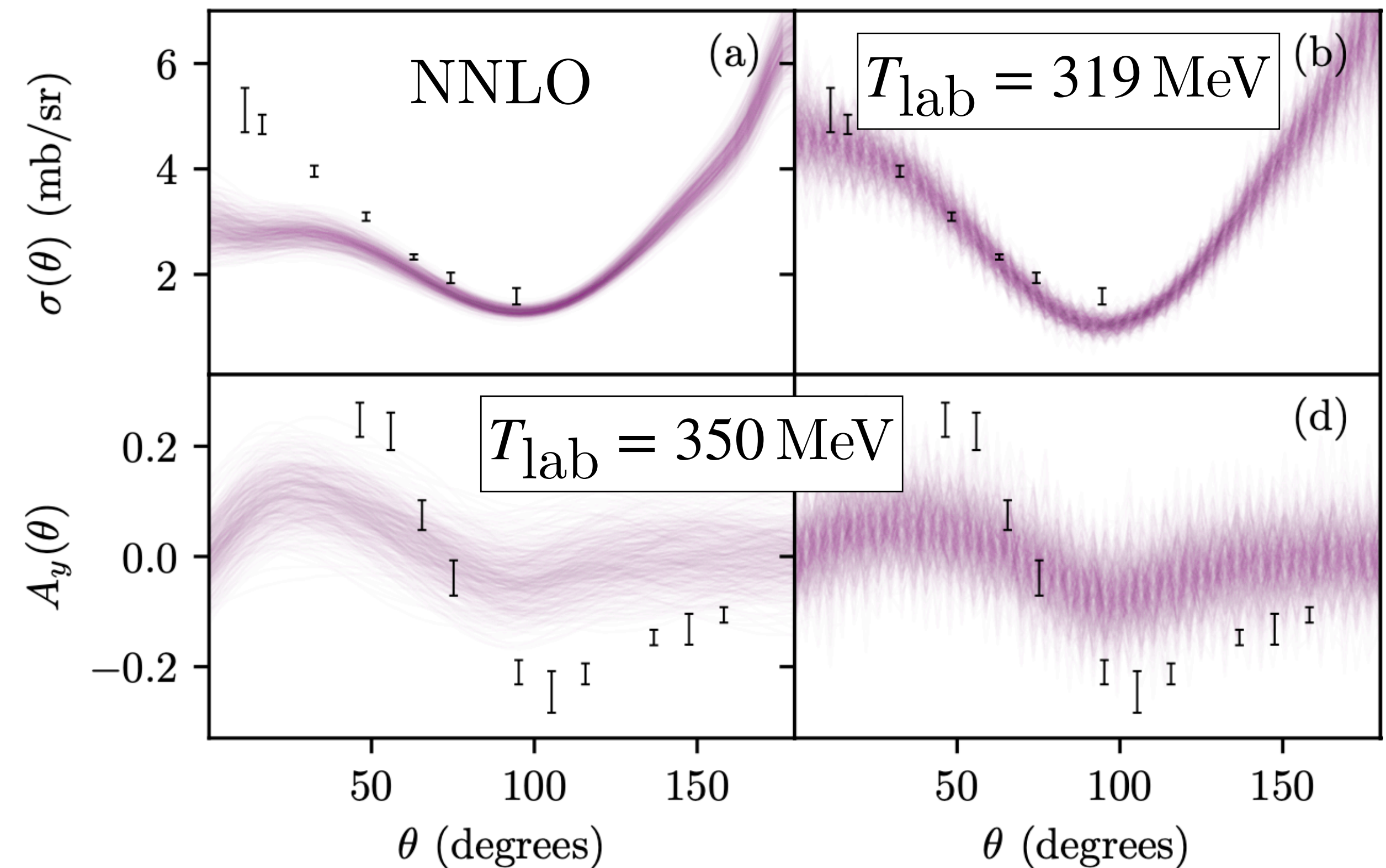
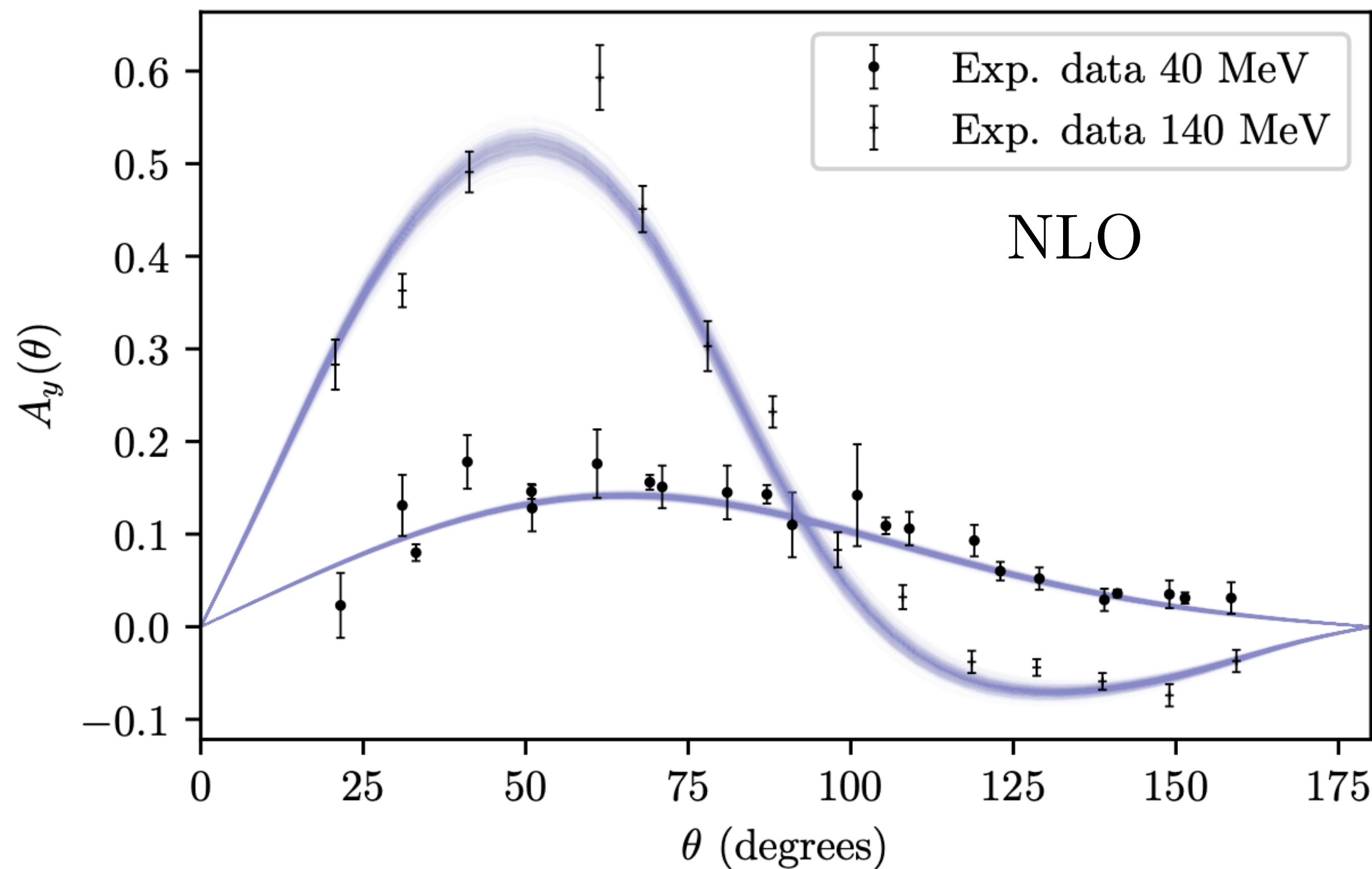
Posterior predictive distributions



seen data

unseen data

(some of the 'bad ones')

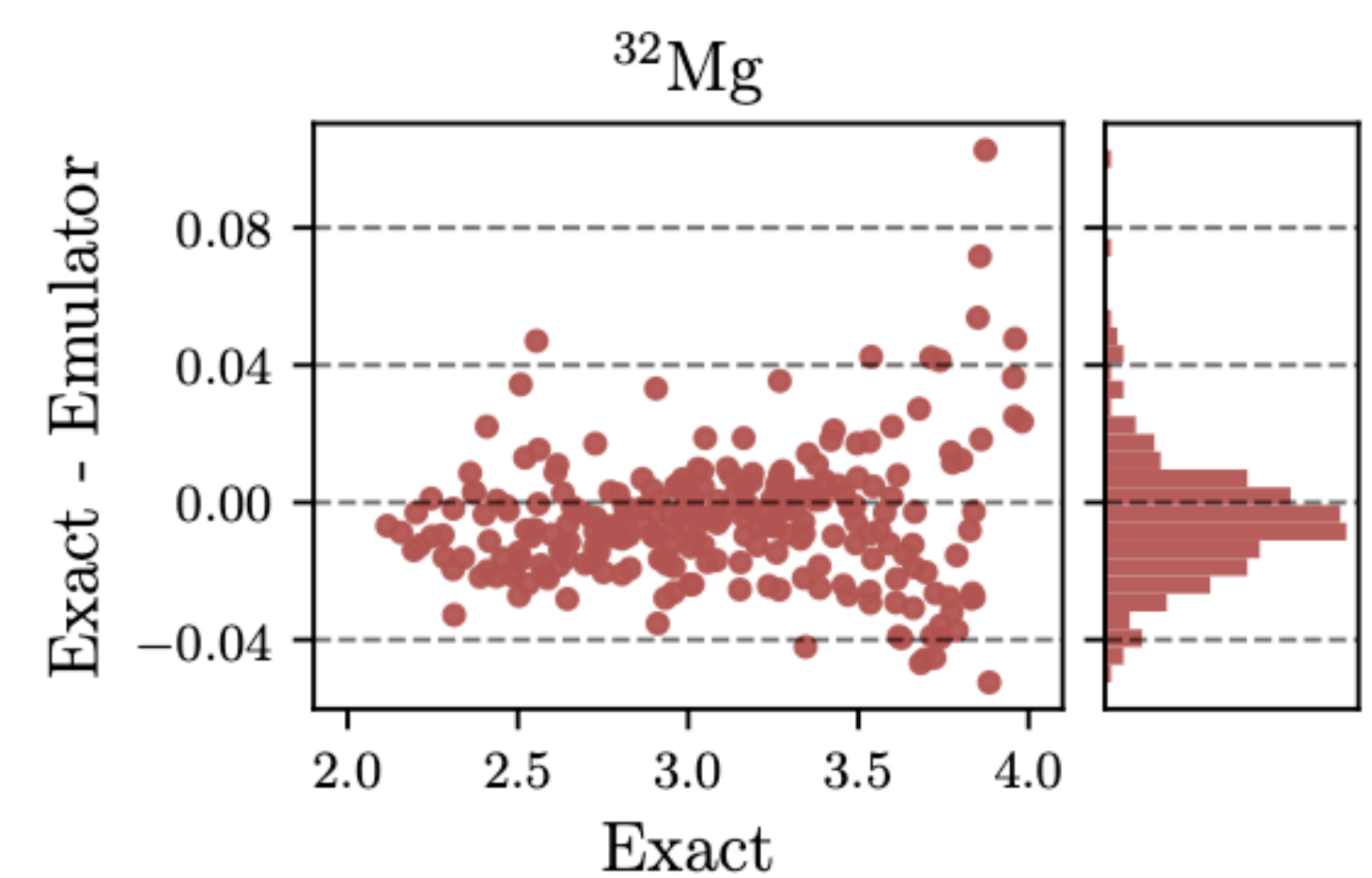
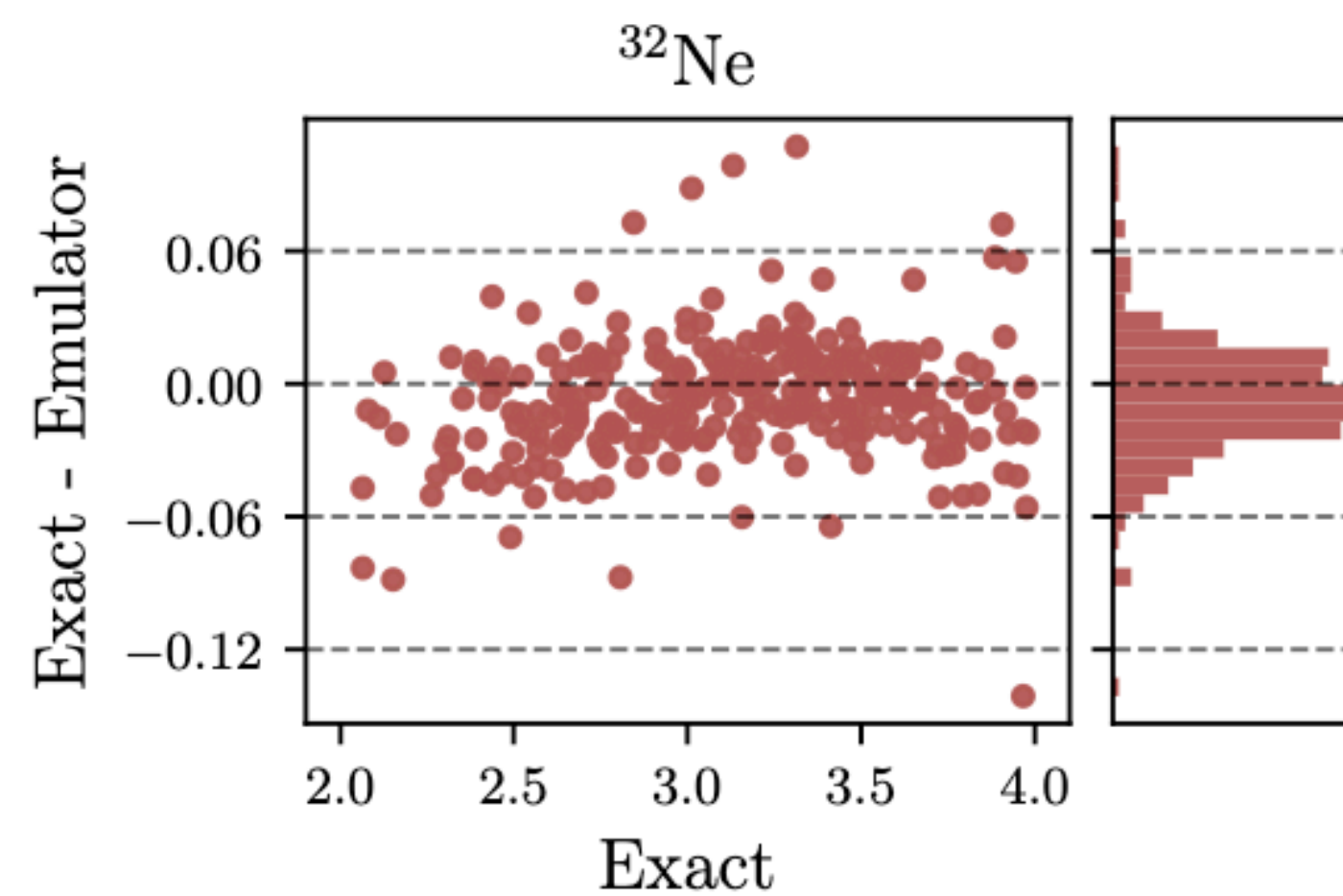
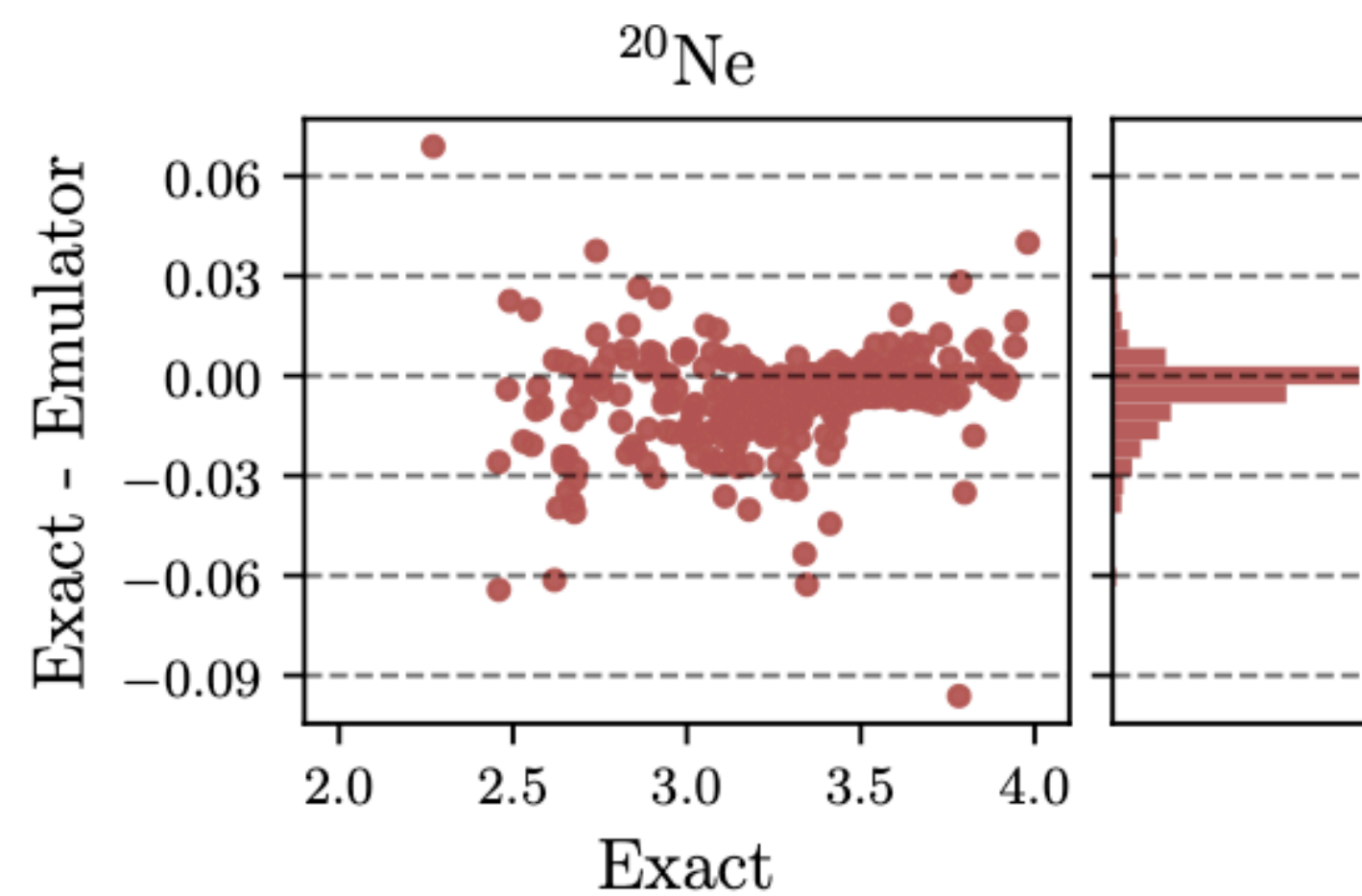
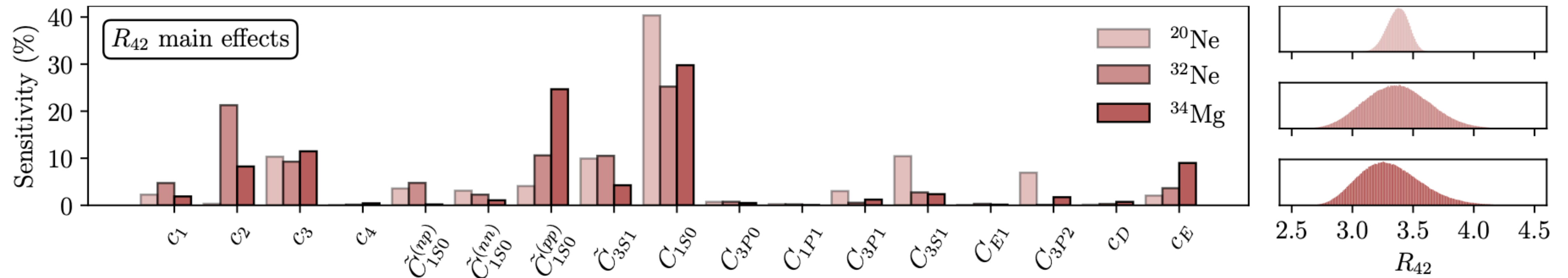


Global sensitivity analysis

Nuclear Few- and Many-Body Problems



Emulating excited state energies of deformed atomic nuclei



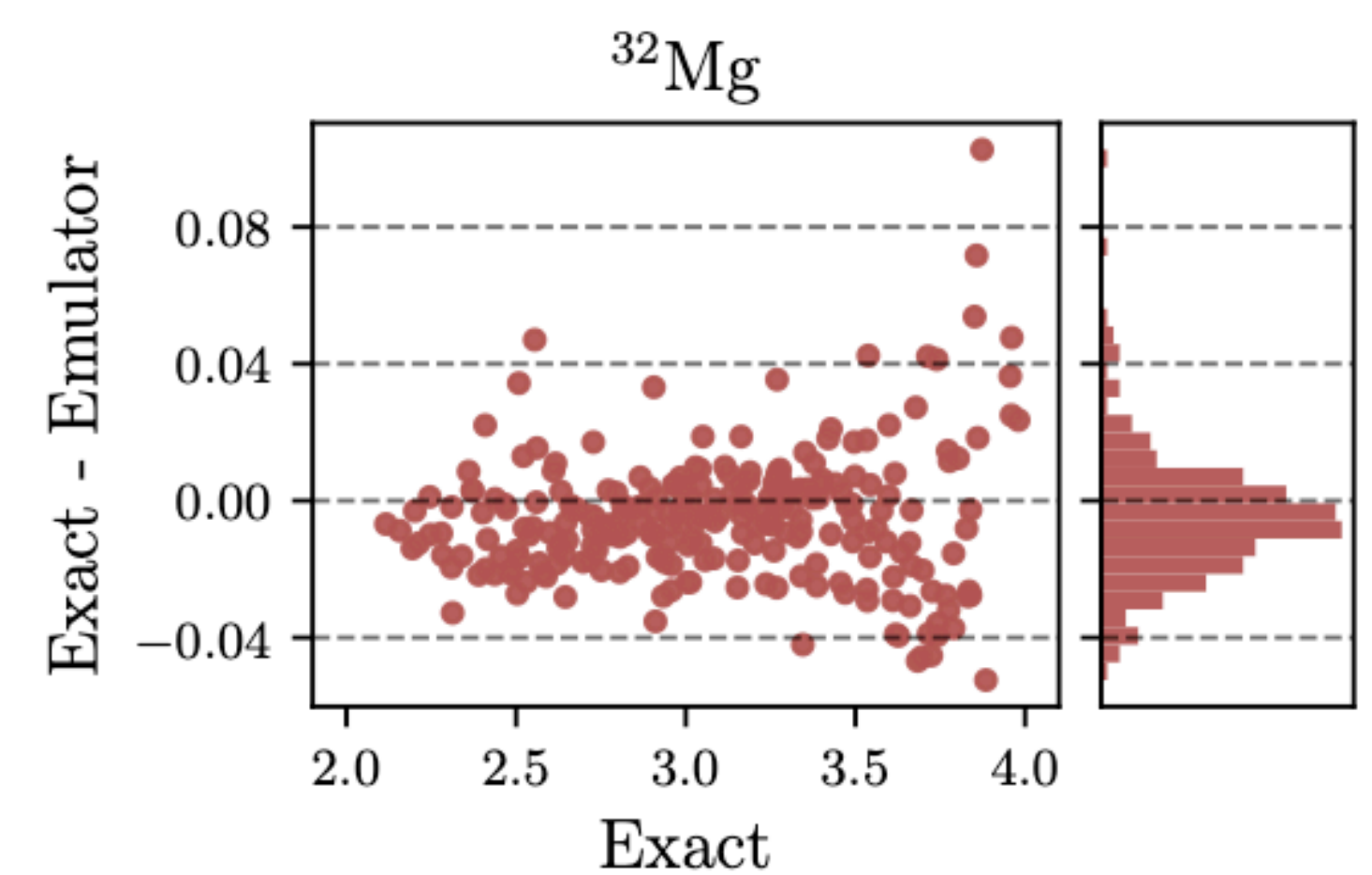
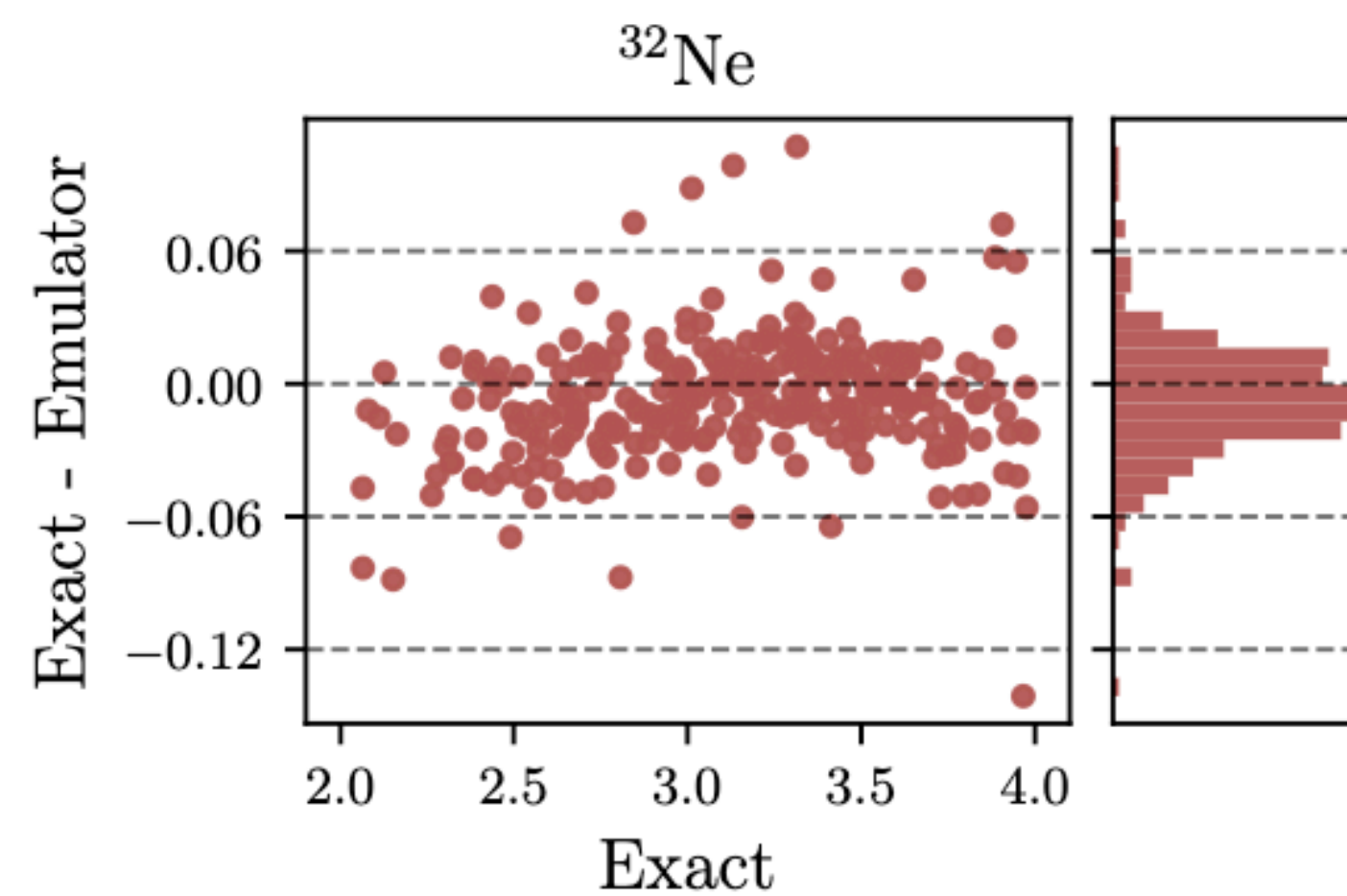
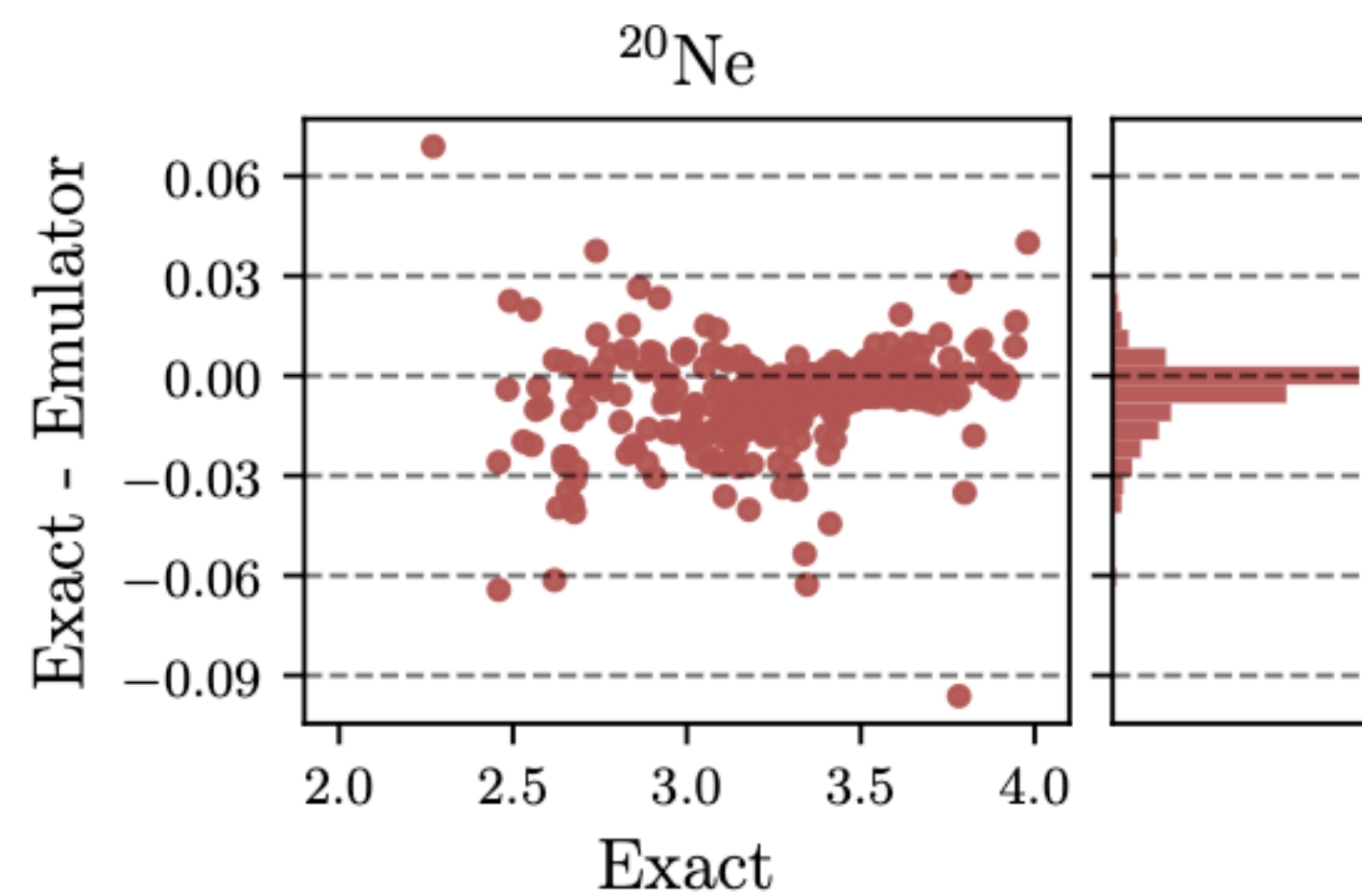
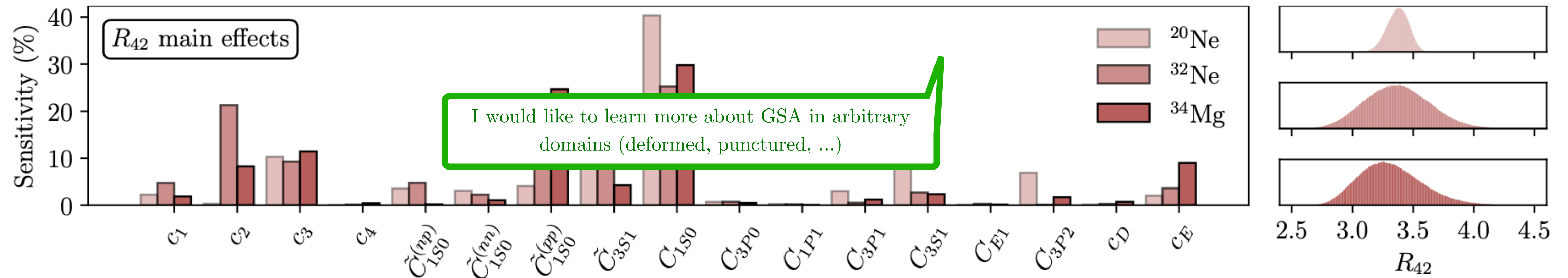
See K. Becker's talk for more on GSA

Global sensitivity analysis

Nuclear Few- and Many-Body Problems



Emulating excited state energies of deformed atomic nuclei



See [K. Becker's talk](#) for more on GSA

- **we developed** a Gaussian process to quantify the correlated model discrepancy (truncation error) of Δ -full chiral effective field theory at NLO and NNLO in np scattering. Much of the essential machinery is in place (open source!).
- **we found** that the *correlated* discrepancy has a small impact in Δ -full EFT up to NNLO. Small difference between NLO and NNLO \rightarrow underestimated prediction errors.
- **problems abound:**
 - tension between πN and NN sectors. (maybe just misspecified discrepancy)
 - sampling in high-dimensional spaces is always challenging.
 - do we have a sensible chiral EFT expansion? (likely not)
 - always a challenge to solve the many-body Schrödinger equation & quantify the associated method errors.
- **emulators** work like a charm and open up for many exciting and useful Bayesian analyses in *ab initio* nuclear physics!

Thanks for your attention!