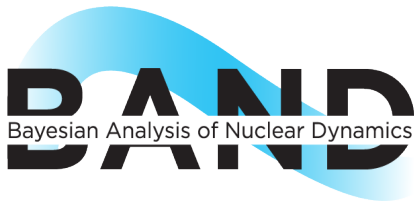


# Sequential Bayesian experimental design for calibration of expensive simulation models

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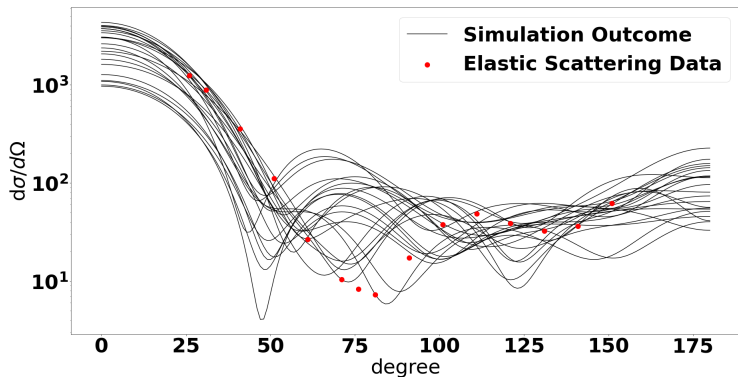
ISNET, May 2023



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# Simulation Models to Understand Complex Processes

- ▶ Requires significant computation time
  - ▶ more complex phenomena require more time-consuming simulations!



- ▶ Let  $(\mathbf{x}_1, \dots, \mathbf{x}_d)$  be the design points where the data  $\mathbf{y} = (y(\mathbf{x}_1), \dots, y(\mathbf{x}_d))$  is collected
- ▶ Seek parameter vector(s)  $\theta \in \mathbb{R}^p$  to align simulator outcomes  $\eta(\theta) = (\eta(\mathbf{x}_1, \theta), \dots, \eta(\mathbf{x}_d, \theta)) \in \mathbb{R}^d$  with observation  $\mathbf{y} \in \mathbb{R}^d$
- ▶ Observation  $\mathbf{y}$  can be modeled using the expensive simulation  $\eta(\theta)$

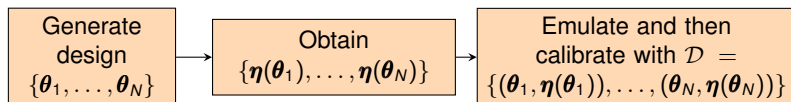
$$\mathbf{y} = \eta(\theta) + \varepsilon, \quad \varepsilon \sim \mathbf{N}(\mathbf{0}, \Sigma)$$

- ▶ Posterior distribution is computed by using Bayes' rule

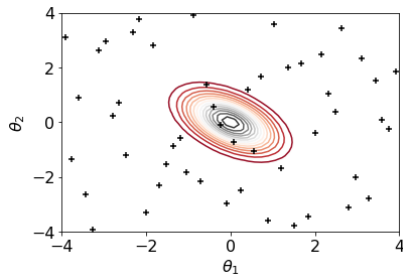
$$\underbrace{p(\theta|\mathbf{y})}_{\text{posterior}} \propto \tilde{p}(\theta|\mathbf{y}) = \underbrace{p(\mathbf{y}|\theta)}_{\text{likelihood}} \underbrace{p(\theta)}_{\text{prior}}$$

- ▶ Prior  $p(\theta)$  can be computed for any  $\theta$  (i.e., independent of simulation)
- ▶ Likelihood requires expensive simulator evaluation for a given  $\theta$

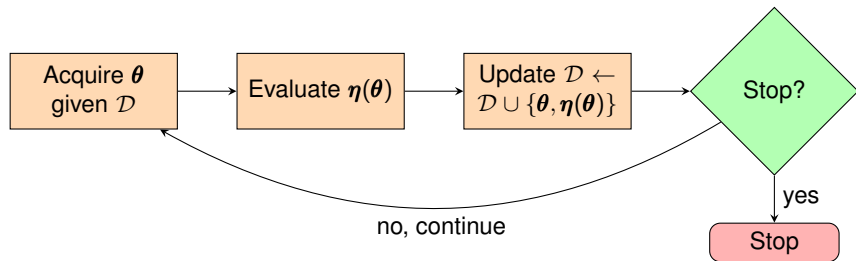
- ▶ MCMC methods require too many evaluations of expensive simulations
- ▶ Alternatively, one builds a cheaper emulator as a proxy to a simulator, and then leverages MCMC sampling to obtain draws from the posterior



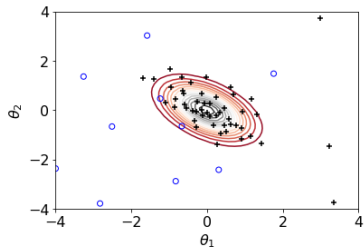
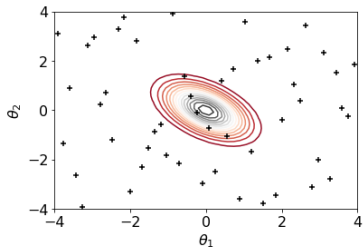
# Example of 50 samples using LHS



# Sequential Bayesian Experimental Design



# LHS vs. proposed approach

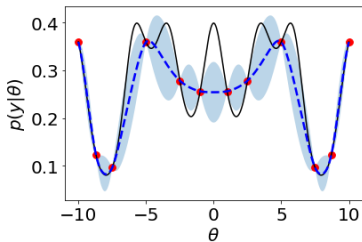
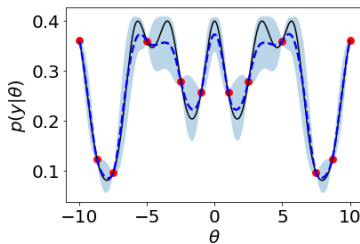
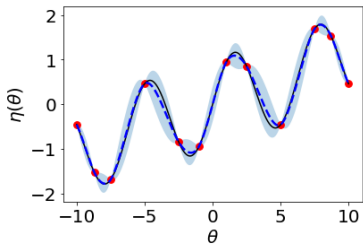


## Algorithm: Sequential Bayesian experimental design

- 1 Initialize  $\mathcal{D}_1 = \{(\boldsymbol{\theta}_i, \boldsymbol{\eta}(\boldsymbol{\theta}_i)) : i = 1, \dots, n_0\}$
- 2 **for**  $t = 1, \dots, n$  **do**
- 3     *Fit* an emulator with  $\mathcal{D}_t$
- 4     *Generate* candidate solutions  $\mathcal{L}_t$
- 5     *Select*  $\boldsymbol{\theta}^{\text{new}} \in \arg \min_{\boldsymbol{\theta}^* \in \mathcal{L}_t} \mathcal{A}_t(\boldsymbol{\theta}^*)$
- 6     *Evaluate*  $\boldsymbol{\eta}(\boldsymbol{\theta}^{\text{new}})$
- 7     *Update*  $\mathcal{D}_{t+1} \leftarrow \mathcal{D}_t \cup (\boldsymbol{\theta}^{\text{new}}, \boldsymbol{\eta}(\boldsymbol{\theta}^{\text{new}}))$



# Emulation of Simulation Output



# Gaussian Process Model for High-Dimensional Simulation Output

$$\Xi_t = \begin{bmatrix} \eta(\mathbf{x}_1, \boldsymbol{\theta}_1) & \cdots & \eta(\mathbf{x}_1, \boldsymbol{\theta}_{n_t}) \\ \eta(\mathbf{x}_2, \boldsymbol{\theta}_1) & \cdots & \eta(\mathbf{x}_2, \boldsymbol{\theta}_{n_t}) \\ \vdots & \vdots & \vdots \\ \eta(\mathbf{x}_d, \boldsymbol{\theta}_1) & \cdots & \eta(\mathbf{x}_d, \boldsymbol{\theta}_{n_t}) \end{bmatrix}$$

GP-based emulator that employs the basis vector approach that is now standard practice (e.g., Higdon et al., 2008)

- ▶ Transform  $d$ -dimensional  $\Xi_t$  into  $q$ -dimensional space
  - ▶  $\mathbf{W}_t = [\mathbf{w}_{t,1}, \dots, \mathbf{w}_{t,q}] = \mathbf{B}_t^\top \Xi_t$
  - ▶  $\mathbf{B}_t = [\mathbf{b}_1, \dots, \mathbf{b}_q]$  stores the orthonormal basis vectors
- ▶ Build  $q$  independent GPs on the latent space

# Gaussian Process Model for High-Dimensional Simulation Output

An independent GP is used for each latent output  $w_j(\cdot)$  with mean  $m_{t,j}(\cdot)$  and variance  $\varsigma_{t,j}^2(\cdot)$  such that

$$w_j(\boldsymbol{\theta}) | \mathbf{w}_{t,j} \sim \mathcal{N} \left( m_{t,j}(\boldsymbol{\theta}), \varsigma_{t,j}^2(\boldsymbol{\theta}) \right), \quad j = 1, \dots, q,$$

where

$$m_{t,j}(\boldsymbol{\theta}) = \mathbf{k}_j(\boldsymbol{\theta}, \boldsymbol{\theta}_{1:n_t}) \mathbf{K}_j(\boldsymbol{\theta}_{1:n_t})^{-1} \mathbf{w}_{t,j}$$
$$\varsigma_{t,j}^2(\boldsymbol{\theta}) = \mathbf{k}_j(\boldsymbol{\theta}, \boldsymbol{\theta}) - \mathbf{k}_j(\boldsymbol{\theta}, \boldsymbol{\theta}_{1:n_t}) \mathbf{K}_j(\boldsymbol{\theta}_{1:n_t})^{-1} \mathbf{k}_j(\boldsymbol{\theta}_{1:n_t}, \boldsymbol{\theta})$$

Transfer predictions back to the original space such that predictive distribution of the emulator output is

$$\boldsymbol{\eta}(\boldsymbol{\theta}) | \mathcal{D}_t \sim \text{MVN}(\boldsymbol{\mu}_t(\boldsymbol{\theta}), \mathbf{S}_t(\boldsymbol{\theta}))$$

# Posterior Inference

Goal is to better learn

$$\tilde{p}(\boldsymbol{\theta}|\mathbf{y}) = p(\mathbf{y}|\boldsymbol{\theta})p(\boldsymbol{\theta})$$

where

$$p(\mathbf{y}|\boldsymbol{\theta}) = (2\pi)^{-d/2}|\boldsymbol{\Sigma}|^{-1/2} \exp\left(-\frac{1}{2}(\mathbf{y} - \boldsymbol{\eta}(\boldsymbol{\theta}))^\top \boldsymbol{\Sigma}^{-1}(\mathbf{y} - \boldsymbol{\eta}(\boldsymbol{\theta}))\right).$$

## Lemma

*Assuming that the covariance matrices  $\boldsymbol{\Sigma}$  and  $\mathbf{S}_t(\boldsymbol{\theta})$  are positive definite,*

$$\begin{aligned}\mathbb{E}[\tilde{p}(\boldsymbol{\theta}|\mathbf{y})|\mathcal{D}_t] &= f_{\mathcal{N}}(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta})) p(\boldsymbol{\theta}), \\ \mathbb{V}[\tilde{p}(\boldsymbol{\theta}|\mathbf{y})|\mathcal{D}_t] &= \left( \frac{1}{2^d \pi^{d/2} |\boldsymbol{\Sigma}|^{1/2}} f_{\mathcal{N}}\left(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \frac{1}{2}\boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta})\right) \right. \\ &\quad \left. - (f_{\mathcal{N}}(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta})))^2 \right) p(\boldsymbol{\theta})^2.\end{aligned}$$

EIVAR criterion is calculated by

$$\mathcal{A}_t(\boldsymbol{\theta}^*) = \int_{\Theta} \mathbb{E}_{\boldsymbol{\eta}^* | \mathcal{D}_t} (\mathbb{V}[\rho(\mathbf{y} | \boldsymbol{\theta}) | (\boldsymbol{\theta}^*, \boldsymbol{\eta}^*) \cup \mathcal{D}_t]) \rho(\boldsymbol{\theta})^2 d\boldsymbol{\theta},$$

where  $\boldsymbol{\eta}^* := \boldsymbol{\eta}(\boldsymbol{\theta}^*)$  represents the new simulation output at  $\boldsymbol{\theta}^*$ .

At each stage  $t$ , the next parameter is chosen to minimize (approximately) the acquisition function  $\mathcal{A}_t$  such that

$$\boldsymbol{\theta}^{\text{new}} \in \arg \min_{\boldsymbol{\theta}^* \in \mathcal{L}_t} \mathcal{A}_t(\boldsymbol{\theta}^*).$$

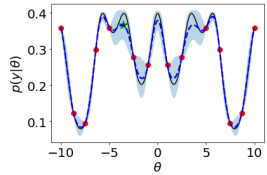
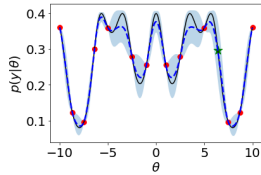
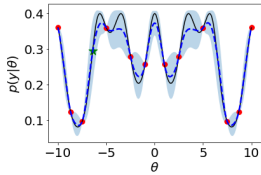
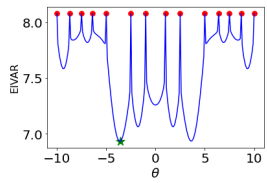
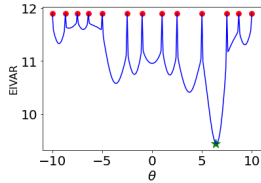
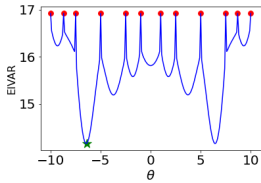
## Lemma

Under the conditions of Lemma 1,

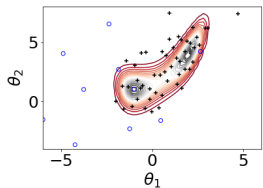
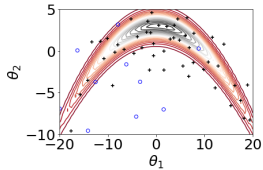
$$\begin{aligned} & \int_{\Theta} \mathbb{E}_{\boldsymbol{\eta}^* | \mathcal{D}_t} (\mathbb{V}[p(\mathbf{y} | \boldsymbol{\theta}) | (\boldsymbol{\theta}^*, \boldsymbol{\eta}^*) \cup \mathcal{D}_t]) \rho(\boldsymbol{\theta})^2 d\boldsymbol{\theta} \\ &= \int_{\Theta} \left( \frac{f_{\mathcal{N}}(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \frac{1}{2} \boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta}))}{2^d \pi^{d/2} |\boldsymbol{\Sigma}|^{1/2}} - \frac{f_{\mathcal{N}}(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \frac{1}{2} (\boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta}) + \boldsymbol{\phi}_t(\boldsymbol{\theta}, \boldsymbol{\theta}^*)))}{2^d \pi^{d/2} |\boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta}) - \boldsymbol{\phi}_t(\boldsymbol{\theta}, \boldsymbol{\theta}^*)|^{1/2}} \right) \rho(\boldsymbol{\theta})^2 d\boldsymbol{\theta}. \end{aligned}$$

Approximate EIVAR is given as

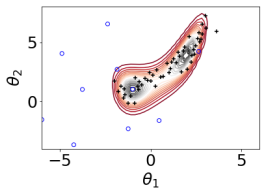
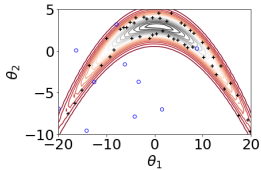
$$\begin{aligned} & \frac{1}{|\Theta_{\text{ref}}|} \sum_{\boldsymbol{\theta} \in \Theta_{\text{ref}}} \rho(\boldsymbol{\theta})^2 \left( \frac{f_{\mathcal{N}}(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \frac{\boldsymbol{\Sigma}}{2} + \mathbf{S}_t(\boldsymbol{\theta}))}{2^d \pi^{d/2} |\boldsymbol{\Sigma}|^{1/2}} - \right. \\ & \left. \frac{f_{\mathcal{N}}(\mathbf{y}; \boldsymbol{\mu}_t(\boldsymbol{\theta}), \frac{1}{2} (\boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta}) + \boldsymbol{\phi}_t(\boldsymbol{\theta}, \boldsymbol{\theta}^*)))}{2^d \pi^{d/2} |\boldsymbol{\Sigma} + \mathbf{S}_t(\boldsymbol{\theta}) - \boldsymbol{\phi}_t(\boldsymbol{\theta}, \boldsymbol{\theta}^*)|^{1/2}} \right). \end{aligned}$$



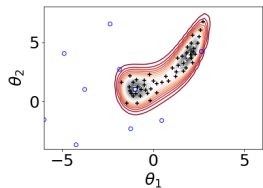
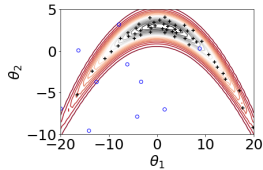
### EIVAR



### MAXVAR

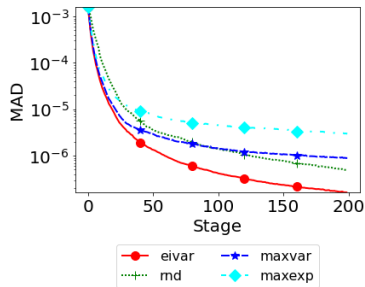
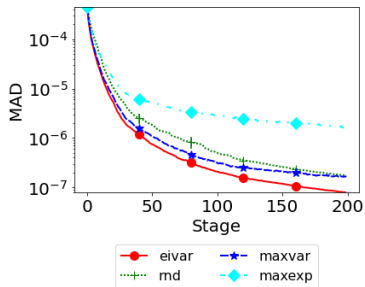


### MAXEXP





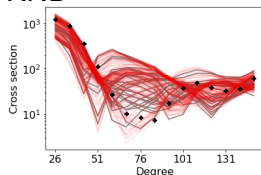
# Comparison of Different Acquisition Functions



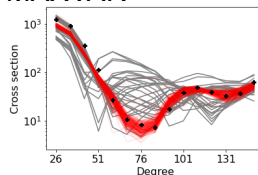
# Application to a Reaction Model

- ▶ Reaction code `FRESCOX` takes the optical model parameters as input and generates the corresponding cross sections across angles from  $0^\circ$  to  $180^\circ$
- ▶ Case study uses the elastic scattering data for the  $48\text{Ca}(n,n)48\text{Ca}$  reaction

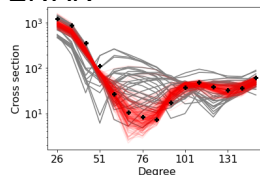
RND



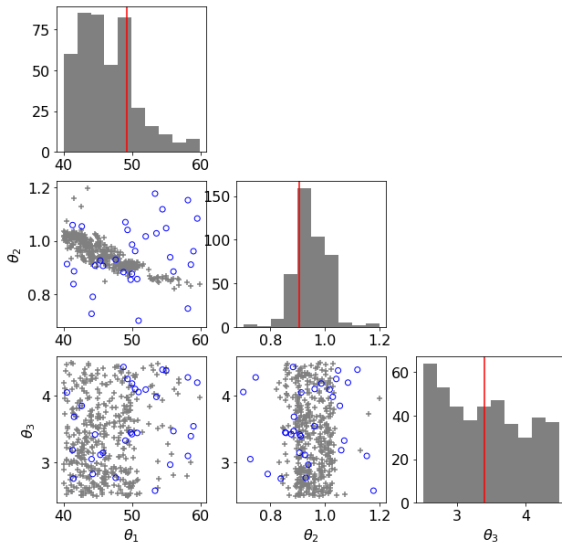
MAXVAR



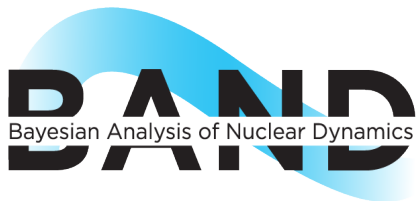
EIVAR



# Application to a Reaction Model



# Questions?



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